# Supplementary material

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## 1. CRF

Figure S1. Content of triage sheet in French. COVID-AMBU study, Switzerland, March-July 2020

Consultation clinique COVID-19	Etiquette patient.e							
Date : / / 2020 ID évaluateur :								
Constantes :       T°      °C       C       ≤35° ou ≥40°C         SpO2      %       C       ≤90%       C	nce Drapeau rouge = Critère d'urgence → Avis médical urgent							
FC / minutes 💭 >125/min 💭 >1 FR / minutes 💭 ≥ 30/min 💭 ≥ 2 TA/mmHg 💭 TAs⊴90 💭 <1 Poids :kg Taille : cm	10/min       Drapeau orange et/ou facteurs de risque =       Consultation médicale         10       → Consultation médicale         En cas de doute, appeler le CDC ou le superviseura de garde							
□ A déjà consulté pour COVID-19. Date : Ancier	n résultat frottis : (pos ou nég.)							
Symptômes :	⊐ T °≥ 38°C —							
Facteurs de risque :         Age ≥ 65 ans         HTA       En présence d'un facteur de risque →         Diabète       Consultation médicale         Maladie cardiovasculaire       En cas de doute, appeler le CDC ou le supervisur de garde         Immunosuppression / □ Cancer       Autre facteur, p.ex. femme enceinte,								
Contact :       □       Exposition professionnelle sans protection à un cas confirmé         □       Contact avec cas suspect ou confirmé vivant sous le même toit alors que la personne était malade ou dans les 24h précédentes         □       Contact étroit (<2 mètres et >15 min) avec un cas confirme	Contexte professionnel : Unisanté CHUV Autre professionnel de la santé Travaille en EMS/ESE (établissement socio-éducatif)							
Orientation :       Fr         □       Frottis diagnostique et retour à domicile avec consignes d'auto-isolement         □       Consultation médicale à la filière COVID-19         □       Consultation aux urgences CHUV	rottis naso-pharyngé ce jour : Effectué par infirmière Effectué par médecin Non effectué ce jour							
Suivi :       Image: Suivi téléphonique à 48h accepté       Image: Suivi téléphonique refusé         Suivi téléphonique refusé       Compare teléphonique refusé	ude COVID AMBU : ormations transmises : Oui ONon onsentement signé : Oui Non Va réfléchir							

Figure S2. Content of triage sheet, translation in English. COVID-AMBU study, Switzerland, March-July 2020

COVID-19 Clinical consultation	Patient							
Date : / / 2020 ID evaluator :								
Impairment of cognitive functions, altered consciousness Body parameters :T° °C $\qquad \qquad \qquad$	Red flags = Emergency criteria → urgent medical notice							
HR/ minutes>125/min>11RR/ minutes $\geq$ 30/min $\geq$ 20BP/mmHgTAs<90<11Weight :kgTaille :cm	0/min 0/min 0 Orange flags and/or risk factors = → Medical consultation							
Has already consulted for COVID-19. Date : Fo	ormer PCR swab result :							
Symptoms : □ Cough □ Dyspnea □ Sore throat □ Fever □ Muscular pain □ Anosmia □ Agueusia / Dysgueusia □ Other								
Risk factors :         □       Age ≥ 65 years olds         □       Arterial hypertension         □       Diabetes         □       Cardiovascular disease         □       Chronic respiratory disease         □       Immunosuppression / □ Cancer         □       Other factor, for example pregnancy								
Contact : <ul> <li>Unprotected occupational exposure to a confirmed case</li> <li>Contact with a suspected or confirmed case living under the same roof when the person was sick or in the previous 24 hours</li> <li>Close contact (&lt;2 meters and &gt;15 min) with a confirmed case</li> </ul> Occupational contexte <ul> <li>Unisanté</li> <li>CHUV</li> <li>Other health professiona</li> <li>Work in socio-educational establishment</li> </ul>								
Orientation :       Diagnostic smear and return home with self-isolation instructions       Nas         Medical consultation to COVID-19 consultations       Adressed to the CHUV emergency	Nasopharyngeal smear of the day : By a nurse By a doctor Not done today							
Follow-up :       CO         Phone follow up at 48h accepted       Info         Phone follow up refused       Co	COVID AMBU study : nformations given : Yes No Consent form signed : Yes No Will think about it							

#### 2. Data management

Data about myalgia and sore throat was only collected after on March 20st, 2020. Information was considered missing before this date.

Education level achieved was grouped into three categories: "High" (University education), "Middle" (Higher secondary education), and "Low" (Lower secondary education or lower).

Occupational position was classified into three categories: "High" (Managers: liberal professions, directors, professors), "Middle" (Lower level executives: teachers, qualified technicians, nurses) and "Low" (Low qualified non-manuals and manuals: sales assistants, clerks, manual workers [22]).

Were categorized as "health workers" professionals working in health structures potentially receiving COVID-19 patients and referring their staff to Unisanté for testing.

## 3. Detailed statistical methods COVID-AMBU study, Switzerland, March-July 2020

### *Calculating the duration of symptom(s)*

The presence or absence of symptoms could only be assessed at the initial consultation and during follow-up visits. Consequently, there was uncertainty regarding the dates at which a symptom started and/or ended, in which case the duration  $D_{ij}$  of symptom j in patient i could not be determined exactly. Rather, it is known to lie within an interval  $L_{ij} \leq D_{ij} \leq U_{ij}$ , with  $L_{ij} \geq 0$  and  $U_{ij} \leq \infty$ . In other words, symptoms duration most often represented either interval-censored (i.e.  $U_{ij} < \infty$ ) or right-censored (i.e.  $U_{ij} = \infty$ .) time-to-event data, except when patients could report the exact dates at which symptoms started and ended.

Let  $s_{ij}$  (start) and  $e_{ij}$  (end) refer to the time at which symptom *j* first appeared and last disappeared in patient *i*, respectively, with j = 1, ..., 5indexing the following five main symptoms exclusively: cough, fever, sore throat, dyspnea and anosmia. Additionally, let  $s_{i0}$  denote the time at which the first symptom (irrespective of which one and also considering symptoms outside of the above list) appeared in patient *i*, and let  $e_{i0}$ denote the time at which that patient became asymptomatic. When these times are available, that is when a patient could report the exact dates symptoms started and ended, it is straightforward to calculate a duration  $D_{ij} = e_{ij} - s_{ij}$  over which patient *i* experienced a given symptom (for *j* = 1,...,5) or a duration over which this patient suffered from at least one symptom (for j = 0). However, this is rarely possible in practice due to missingness in either  $s_{ij}$ ,  $e_{ij}$  or both. Such missingness induces uncertainty in  $D_{ij}$ , which is then no longer observable exactly. Since the presence or absence of symptom(s) could only be assessed during followup calls, all we know rather is that the duration of symptom(s) lies within an interval of time, such that  $L_{ij} \leq D_{ij} \leq U_{ij}$ , with the lower and upper bound of that interval (i.e.  $L_{ii} \ge 0$  resp.  $U_{ii} \le \infty$ ) being known exactly.

Let us first consider individual symptoms i.e. j > 0 and denote by  $t_{ik}$ the date of the *k*-th follow-up call to patient *i* (with k = 0 referring to the initial consultation). Suppose that symptom *j* appeared between e.g. the first and second follow-up calls for this patient, so that  $t_{i1} < s_{ij} < t_{i2}$ . Additionally, suppose that symptom *j* disappeared e.g. between the third and fourth follow-up calls, so that  $t_{i3} < e_{ij} < t_{i4}$ . In that example, the duration  $D_{ij}$  is interval-censored between  $L_{ij} = t_{i3} - t_{i2}$  and  $U_{ij} = t_{i4} - t_{i1}$ . When symptom *j* was already present at the initial consultation,  $s_{ij}$  was considered interval-censored with  $s_{i0} \le s_{ij} < t_{i0}$ . However, when  $s_{i0}$  was missing, that is when the patient could not recall when he/she first experienced the first symptom, we assumed that the first symptom appeared within the 14 days preceding the initial consultation i.e.  $t_{i0}$  –  $14 < s_{i0} < t_{i0}$ . A delay of 2 weeks before the initial consultation was chosen because the first symptom appeared within about that timeframe in 90% of the patients with non-missing  $s_{i0}$  (for 75% of these patients, the first symptom appeared within 7 days before the consultation). Similarly, when symptom *j* was still present during the last follow-up call to patient *i* (say at  $t_{i4} = t_{i0} + 28$  days), we assumed that it disappeared between that last follow-up and the date at which the patient became asymptomatic, such that  $t_{i4} < e_{ij} < e_{i0}$ . Such situations would all lead to  $D_{ij}$  being interval-censored. However, when  $e_{i0}$  was missing, that is when the patient could not report the exact date at which the last symptom disappeared, we only considered  $e_{ij} > t_{i4}$  and consequently treated  $D_{ij}$  as being right-censored (i.e.  $U_{ij} = \infty$ ). The duration  $D_{i0}$  with at least one symptom was treated similarly but unlike for  $D_{ij}$  with j > 0 which is either interval- or right-censored, we note that  $D_{i0}$  could sometimes be observed exactly when both  $s_{i0}$  and  $e_{i0}$  were reported.

Furthermore, note that a patient was assumed to experience a given symptom continuously between the first and the last reported occurrence of that symptom. Temporary disappearance of a symptom was not considered as relevant and the severity of the symptom was not taken into account in the analysis.

## Multiple imputations

Multiple imputations (MI) [2] were used to impute  $D_{ij}$  in the interval  $[L_{ij}; U_{ij}]$  when  $U_{ij} < \infty$ . They were also used to impute missing covariates values (including missing RT-PCR results for patients who were not tested). A total of 30 imputed datasets was constructed.

Given a set of variables, the technique of multiple imputations (MI) [23] proceeds by modelling each variable with missing values in turn as a function of the other variables. Missing values in the outcome variable are imputed given other covariate values and the process is repeated after selecting another variable as the outcome. Each missing value is imputed *M* times in order to build *M* complete datasets which can then be analyzed. In this study, we used M = 30 datasets. At last, results from these *M* analyses can be pooled using Rubin's rules [23].

The duration  $D_{ij}$  was modelled using a Cox proportional hazard regression, adjusting for several covariates (see Supplementary Table 1 for a comprehensive list) including durations of other symptoms (i.e.  $D_{ik}$  for all  $k \neq j$ ). When including a duration as a predictor in an imputation model, we followed the approach of White [24] and included both the cumulative baseline hazard (as calculated with the Nelson-Aalen estimator) and the event indicator (0=right-censored, 1=observed or interval-censored). Additionally, we also included a binary indicator  $I_{ij}$ taking the value 1 if patient *i* did not experience symptom *j* at all, and 0 otherwise. This was necessary because when  $I_{ii} = 1$ , both the cumulative baseline hazard and the event indicator for symptom *j* are undefined had to be set to zero. Continuous variables (e.g. BMI, body temperature) were modelled using linear regression, dichotomous variables (e.g. PCR result, presence of  $\geq 1$  risk factor) were modelled using logistic regression, and ordinal variables (highest education level and job activity) were modelled using cumulative logit regression. Note that both age and BMI did enter the imputation models as continuous predictors but they were later categorized in the multivariate Cox regression analysis (Table S1).

At iteration 0 of the imputation algorithm, any interval-censored duration  $D_{ij}$  was imputed by randomly drawing a value within the interval  $[L_{ij}; U_{ij}]$  according to a uniform distribution. This was performed separately in each of the *M* imputed datasets. Imputed durations were then treated as "observed" values throughout the whole imputation process. At subsequent iterations, any interval-censored duration  $D_{ij}$  was randomly imputed in the interval  $[L_{ij}; U_{ij}]$  according to the survival distribution estimated by the Cox proportional hazard model fitted at the preceding iteration. Note that only interval-censored durations required imputation and were then treated as "observed" values. Right-censored

durations were left untouched since Cox regression models already handle such data appropriately. The imputation process was cycled over 50 iterations. Convergence of the imputation algorithm was confirmed by monitoring the mean and standard deviation of each imputed variables over the iterations and ensuring a proper mixing of these parameters across the *M* imputed datasets.

We followed the procedure used in [25] and described in [26] to combine KM curves following MI. Namely, KM estimates were first calculated at predefined time points (every day, ranging from 0 to 60 days) in order to make them comparable across imputations. A complementary log-log transformation [27] was then applied to the daily KM estimates obtained in each imputed dataset before pooling the results using Rubin's rules [23]. The statistical significance of the difference between KM curves obtained for negative and positive RT-PCR results was assessed using the log-rank test, after pooling the test statistics (chi-squared) obtained on each imputed dataset using the method described in [28].

Covariates included in the model were selected a priori based on clinical relevance and are detailed in table S1.

Variable	Туре	Missing values		
Consultation center	Categorical	none		
Age at initial consultation	Continuous	none		
Gender	Binary: male/female	none		
BMI at initial consultation	Continuous	2.8%		
Health care professional	Binary: yes/no	none		
Presence of $\geq 1$ risk factor(s)*	Binary: yes/no	18.5%		
Smoker	Binary: yes/no	6.1%		
Use of AINS in last 7 days prior to consultation	Binary: yes/no	8.6%		
Professional exposition with confirmed case	Binary: yes/no	none		
Contact with a suspected/confirmed case living in the same household	Binary: yes/no	none		
Close contact with a confirmed case	Binary: yes/no	none		
Highest education level	Ordinal: low/mid/high	3.9%		
Type of job activity	Ordinal: low/mid/high	7.1%		
PCR result (≥ 1 positive result during follow-up period)	Binary: neg./pos.	17.4%		
Body temperature at initial consultation	Continuous	6.0%		
Cardiac frequency at initial consultation	Continuous	7.1%		
Consultation period (after Apr. 27, 2020)	Binary: 0/1	none		
Indicator $I_{ij}$ for absence of symptom $j$ in patient $i$	Binary: 0/1	none		

**Table S1.** Covariates entered in imputation models with their type and proportion of missing values. COVID-AMBU study, March-July 2020

Cumulative hazard for $D_{ij}$ (set to 0 if $I_{ij} = 1$ )	Continuous	none
Event indicator for $D_{ij}$ (set to 0 if $I_{ij} = 1$ )	Binary: 0/1	none

Note that the hazard here refers to the event "disappearance of symptom(s)". Consequently, a hazard ratio larger than one corresponds to an increased "risk" of experiencing the event, which in turns corresponds to a shorter duration of symptom(s). Conversely, a hazard ratio smaller than one corresponds to a longer duration of symptom(s).

Missi Symptoms and signs data		Total (N=883)		RT-PCR positive (N=123)		RT-PCR negative (N=606)		Not tested (N=154)		p-value of PCR-pos. vs neg.	
		n	%	n	%	n	%	n	%		
Cough	3	607	(69.0)	99	(80.5)	398	(65.9)	110	(71.9)	0.002	
History of fever	5	436	(49.7)	80	(65.6)	292	(48.4)	64	(41.8)	0.001	
Sore throat	72	434	(53.5)	57	(51.8)	322	(55.3)	55	(46.2)	0.498	
Myalgia	98	418	(53.3)	77	(68.8)	282	(50.5)	59	(51.3)	< 0.001	
Dyspnea	6	341	(38.9)	49	(40.2)	236	(39.2)	56	(36.6)	0.843	
Headache	224	279	(42.3)	52	(49.5)	185	(42.9)	42	(34.2)	0.222	
Fatigue	257	221	(35.3)	41	(40.6)	153	(38.0)	27	(22.1)	0.627	
History of temperature ≥ 38°C	5	193	(22.0)	41	(33.9)	112	(18.5)	40	(26.1)	< 0.001	
Rhinorrhea	225	191	(29.0)	33	(32.0)	134	(30.9)	24	(19.8)	0.819	
Chest pain	259	165	(26.4)	20	(20.0)	119	(29.3)	26	(22.0)	0.062	
Hypo-/a-gueusia	190	148	(21.4)	44	(51.2)	71	(13.9)	33	(34.4)	< 0.001	
Hypo-/a-nosmia	185	135	(19.3)	41	(47.1)	60	(11.7)	34	(34.0)	< 0.001	
Digestive symptoms <sup>1</sup>	236	132	(20.4)	23	(21.9)	95	(22.4)	14	(12.0)	0.921	
Chills	278	73	(12.1)	10	(10.5)	53	(13.6)	10	(8.8)	0.455	

**Table S2.** Symptoms and signs reported by COVID-19 suspect patients, by SARS-CoV-2 RT-PCR testing and test results. COVID-AMBU study,March-July 2020

Abdominal pain	284	69	(11.5)	8	(5.2)	52	(13.5)	9	(7.8)	0.152
Dyspnea > 4 days	6	63	(7.2)	11	(9.0)	45	(7.5)	7	(4.6)	0.561
Fever > 4 days	6	34	(3.9)	10	(8.3)	17	(2.8)	7	(4.6)	0.004
Sweating	307	33	(5.7)	9	(9.5)	20	(5.4)	4	(3.5)	0.147
Signs mean (SD)										
Temperature in °C	53	36.4	(0.6)	36.7	(0.8)	36.4	(0.6)	36.3	(0.6)	0.014
Oxygen saturation, In %, mean (SD)	48	97.1	(1.5)	96.7	(2.3)	97.2	(1.2)	97.0	(1.5)	0.001
Respiratory rate, per min, mean (SD)	315	17.6	(5.5)	20.4	(10.4)	17.1	(3.9)	18.7	(4.8)	< 0.001
Heart rate, per min, mean (SD)	63	84.1	(14.9)	87.7	(16.0)	83.4	(14.6)	84.1	(14.1)	0.006
Systolic pressure, in mmHg, mean (SD)	208	125	(18.0)	126	(18.0)	125	(18.0)	123	(16.0)	0.803
Diastolic pressure, in mmHg, mean (SD)	208	81	(12.0)	83	(14.0)	81	(12.0)	81	(12.0)	0.154
BMI in kg/m2 (SD)	25.0	25.2	(4.9)	26.2	(5.0)	25.1	(5.1)	24 7	(4.0)	0.035

<sup>1</sup>Digestive symptoms: reporting of nausea, vomiting or diarrhea