# MAJOR POWER CAPABILITIES AND INTERSTATE WARS:

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#### INTRODUCTION

Since the end of World War II, the international system (a system that comprises all nations in the world) has experienced a decrease in the occurrence of international wars accompanied by a steady increase of intrastate wars (also referred to as civil war, civil conflict), conflicts that occur within a nation-state rather than between two or more nation states. Many scholars have attempted to identify the characteristics that lead to civil war onset (Anyanwu, 2003; Fearon and Laitin, 2003; Gurr, 1968; Lake, 2003; Mansfield and Snyder; Posen, 1993 2002; Sambonis, 2001; Saxton, 2005; Tilly, 2003) by mainly focusing on the characteristics of countries and how these either promote or preclude the occurrence of a civil war. However, the effects of systemic factors (the aggregate characteristics of the international system rather than characteristics of its members individually) on civil war onset remains impressively understudied. Additionally, scholars have overlooked the potential causes of the frequency of civil wars within the international system over time. In order to fill these gaps, this study aims at answering the following question: Does the hegemon's level of capabilities impact the number of civil wars occurring within the international system? The international system represent the world under which countries operate and interact; the hegemon represents the stronger state in the said system - in this paper, the hegemenon is the United States for the entire time period and the capabilities of the hegemon represent the portion of total world power (the aggregated international power) that the hegemon owns in terms of electricity consumption, military power, and population indices. The paper is organized in the following manner: first, the data sources are identified and explained; second, some background on times series variables and processes are defined; then, the characteristics of the data are presented; the fourth section presents the statistical analyses and the interpretation thereof; the final questions addresses implications, conclusions and avenues for future research.

DATA

In this study, I propose that the capabilities of the main power in the international system represent a leading variable of the frequency of interstate wars over time. Thus, the degree of capabilities is my independent variable and frequency of interstate war is the dependent variable. I utilize the Correlates of War (COW) data on National Material Capabilities and use the data for the US since it has been identified as the hegemon for the period under study – 1946 through 2001 (Singer *et al.*, 1972; Singer, 1987 – the original dataset and the subsequent revisions remain updated through the COW project on an ongoing basis). As the official website indicates, "[t]he National Material Capabilities data set contains annual values for total population, urban population, iron and steel production, energy consumption, military personnel, and military expenditure of all state members, currently from 1816-2001. The widely-used Composite Index of National Capability (CINC) is based on these six variables and included in the data set." This dataset is accessible online (available at http://cow2.la.psu.edu/). I use the CINC as an indicator of US capabilities; this indicator is a fraction that represents the proportion of total international capability possessed by a given country. For the US, this variable takes values ranging from 0.131 to 0.364.

In order to account for the number of interstate wars occurring within the international system each year, I use the Uppsala Armed Conflict Dataset on civil conflict (Eriksson *et al.*, 2003; Gleditsch *et al.*, 2002). This research group constructed a "Monadic Table" that presents civil unrest occurring in every country each year; additionally, each observation contains a "count" column that adds all civil unrest for each country per year. Of relevance to this study are

the last two types of conflict they identify, mainly "internal armed conflict" and "internationalized internal armed conflict"; the former represents wars between a government and an opposition group while the second refers to the same phenomenon with the addition that the opposition is backed by a foreign government. The Uppsala project classified an event as a "war" when a country reaches at least 25 conflict-related civilian deaths within any given year. Because this project focuses on the effect of a systemic component (the share of total capabilities owned by the hegemon), the aggregated value of this variable represents the dependent variable. As such, the "war" series contains the yearly summation of civil wars in the international system. This variable takes values that range from 14 to 81. I can now start looking at the characteristics of both variables in terms of stationarity.

### TIMES SERIES STATISTICS AND MODELS

Unlike traditional cross-sectional data, time series variables contain several unique dynamics that a researcher needs to identify in order to adequately come up with a model that addresses time-series characteristics. When faced with this kind of data, the investigator first needs to determine whether a series is stationary of not. A stationary process is a stochastic process whose probability distribution at fixed time or position is the same for all times or positions; as such, the unconditional variance and mean of such a process remains constant at different points in time. Furthermore, a variable is non-stationary, it usually contains a unit root whereby one or more of the coefficients in an autoregressive model of order 1 (explained below) has a value superior or equal to one (for other types of autoregressive models, one needs to conduct a unit root test to decipher whether a series is stationary or not). If a unit root is present in the model, the latter has a stochastic trend and is integrated – denoted as I(0), I(1), I(2), etc.

variable to the current one for one difference) in order to be stationary; an I(0) series denotes a variable that is stationary "as is", an I(1) necessitate one difference in order to be stationary, and so on. As a result, a stationary variable represents the sole type of variable for which one can estimate a reasonable model. However, many time series data have non-stationary characteristics insofar as they may have time-dependent heteroskedasticity (non-constant variance overtime) an aspect which necessitates some manipulation of the variable in question in order to render it stationary. Usually, differencing the original version of a non-stationary variable suffice to make it stationary. There exist several tools and techniques to decipher whether a series is stationary or not.

The first step in identifying whether a variable is stationary or not consists in graphing it over time. When graphing a series, one can decipher whether a variable is stationary or not. If the data appears to have a constant variance and mean over time, then it most likely is stationary. On the other hand, if the data behaves in an unpredictable manner (it either appears to have an inconstant variance over time, or a time trend, or both), then mist likely is non-stationary. Such a variable can be described as a "random walk", of which there are three types. A "simple random walk" depicts a variable that has no intercept and has non-constant variance over time. A "random walk with drift" has an intercept and shares the other characteristics of a "simple random walk". Finally, a "random walk with time trend and drift" has an intercept and also some sort of time-dependent variation: the values of the variable either accrue or decrease across time. Once a series has been examined from a graphical perspective, the researcher conducts further tests to decide whether the series is indeed stationary of not from a statistical perspective, and if so, what sort of steps need be undertaken. The first statistical tool available consists of looking at the autocorrelation plots (ACF) of the series. This plot simply shows autocorrelations for data values at different points in time: it presents the different lags of the series and presents the levels to which the values of the series at time t are correlated to previous values. If the ACF quickly declines to zero, it indicates that the series most likely mist likely is stationary. However, if it really slowly declines to zero, it signifies that the series is non-stationary. The partial autocorrelation of a value at lag k shows the correlation between the variable at time t and t-k that is not accounted for by lags 1 through lag k-1. Additionally to helping identifying the kind of model needed for a series, the PACF also help finding the number of augmentations needed for a Dickey-Fuller test (subsequently DF or ADF for Augmented Dickey-Fuller test).

Once the ACF and PACF of a series have been studied and identified, the next step consists in conducting the DF. A DF tests whether a unit root is present in an autoregressive model (on AR process may have unit roots). The ADF is an improved version of the original test that deals with more complicated time series variables (this is the test used in this paper). There exist three types of test: a single mean test, an intercept and mean test, and a time trend test. Those three tests directly relate to the three different types of random walks, therefore, the type of test to use is based on the plot of the data. Additionally, the test has several possible augmentations. The appropriate number of augmentations is determined by the PACF: one looks at the number of significant lags in the PACF and uses that at the number of needed augmentations. In this paper, the null hypothesis states that he variable is non-stationary. Thus, if one fails to reject the null, one concludes that the series is non-stationary. If the series is non-stationary, it needs be differenced (by basically subtracting the current value by the previous value, losing one observation in the data) and tested again for stationarity. Several differences

may be needed though most time series variable are integrated of order 1 (or I(1)), meaning that they need one difference to become stationary. After making the series stationary, the researcher attempts to select the "best-fitting" model for the series.

The primary tools for deciding whether a series necessitates a WN, AR, or MA process are the ACFs and PACFs. A WN process has "no flavor" and has an ACF and PACF that never is statistically significant. The equation form of such a model is  $Y_t = \varepsilon_t$  and such a model cannot be estimated for the behavior of the series is totally unpredictable. An AR process has an ACF function that gradually, but relatively quickly, declines to zero. On the other hand, its PACF function is significant for p lags. The value of p helps determine the order of the AR process, i.e., the number of lagged values of the series that should have a statistical significance. Such a model simply tests the correlation between a variable and its past values. The equation form of these models is:  $Y_t = \beta_0 + \beta_1 Y_{t-1} + ... + \beta_p Y_{t-p} + \varepsilon_t$ . An MA process has an ACF that is significant for q lags and a PACF that gradually, but relatively quickly, declines to zero. The value of qdetermines the order of the MA. Such a process states that the values of the dependent variable depend on shocks in the past:  $Y_t = \theta_0 + \epsilon_t - \theta_1 \epsilon_{t-1} - \dots - \theta_q \epsilon_{t-q}$ . Additionally, a variable may contain both AR and MA components and is thus said to an ARMA process. The ACF and PACF function cannot help determine whether a variable necessitates an ARMA process. Thus, in order to decipher whether a model is AR, MA, or ARMA, it is helpful to use information based criterion such as an Aikake Information Criteria (AIC) by estimating different models and figuring out which has the best AIC. However, on theoretical grounds, this study proposes to elucidate on the potential relationship between two variables rather than doing a simple AR, MA, or ARMA on a variable of interest. Consequently, the above-mentioned models are inadequate and a transfer function is more appropriate.

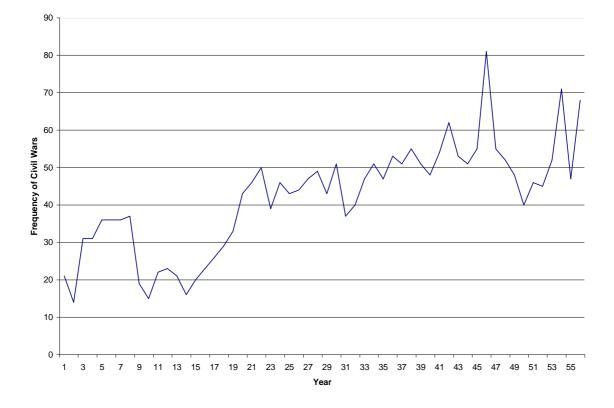
A transfer function takes the following form:  $Y_t = \beta_0 + \beta_1 X_t + ... + \beta_p X_{t,p} + \varepsilon_t$ , where *p* relates to the last lagged dependent variable included in the model. X is the leading indicator of Y: changing values of X overtime account for changing values of Y overtime. Ideally, theory should help determine whether one wants to use a transfer function or not. Unfortunately, even good theory may postulate that a there exist a relationship between X and Y – whereby X supposedly causes Y – statistics may prove otherwise. In order to determine whether a posited transfer function is appropriate, one looks at the cross-correlation of the two proposed variables. Cross-correlations help measure the extent to which two series are related. The cross-correlation functions shown by SAS show both negative and positive lags; if the negative lags are significant, the two variables are not fit for a transfer function. On the other hand, if only positive lags are significant, then, the variables should fit a transfer model. Additionally, the number for which the lags are significant determine the number of lagged values of X that need to be included in the model. Now that the specifics of time-series data and the different types of potentially applicable models here have been identify, it ensues that the first task consists in testing for stationarity for both variables.

#### DATA CHARACTERISTICS

The first necessary step entails identifying the characteristics of the dependent variable, i.e., the frequency of civil war (all commands and log window outputs are included in the Appendix). Figure 1 show the plot for this variable. Evidently, it appears that the number of civil war from 1946 to 2001 has a time trend and an intercept. This preliminary finding seems to indicate that the "civil wars" series contains heteroskedastic components but the finding remains too weak to ascertain that the series is non-stationary, necessitating further tests. These tests will help elucidate whether the variable is stationary or not and whether it needs being differenced in

order to become stationary. As explained in the previous section, the next step involves deriving the ACF and PACF for the non-differenced series; then, the appropriate ADF unit root test is identified.





Graph 1 shows the behavior of the "civil wars" ACF. The series gradually, and rather quickly, declines to zero, which could be an indicator of an AR process or, alternatively, of a non-stationary series.

Graph 1: Autocorrelation Function for Frequency of Civil Wars

Lag	Covariance	Correlation	-19876543	2 1 0 1 2 3 4 5 6 7 8 9 3	1 Std Error
1	151.585	0.74372	.	*********	0.133631
2	137.010	0.67221		*******	0.193936
3	116.320	0.57070	.	*******	0.231839
4	104.089	0.51069		*******	0.255698
5	98.554733	0.48354		*********	0.273306
6	88.252414	0.43299	.	*******	0.288177
7	91.578922	0.44931		*******	0.299570
8	104.221	0.51134	.	********	0.311371
9	79.657958	0.39082	.	*******	0.326022

10	73.952578	0.36283		******	0.334283
11	52.511867	0.25764	Ì	*****	0.341243
12	56.033528	0.27492		*****	0.344699
13	46.541158	0.22834	Ì	*****	0.348593
14	38.872130	0.19072		****	0.351254
				1	

In order to figure out which type of unit root test is required, the level of significance of the PACF indicates this. This paper already demonstrated that the series has both an intercept and a time trend. The PACF presented in Graph 2 only contains one significant lag (mainly, the first one), therefore, we need to check whether the Augmented Dickey Fuller unit root test is significant for one lag with intercept and unit trend. The unit root test with these specifications has a degree of significance of .0976. Normally, one would want a 95% confidence in this number; however, due to the limited amount of data, this value may suffice for the purpose of this paper. Further stationarity tests over the dependent variable will help decide whether it needs to be differenced in order to be stationary.

Lag	Correlation	-1 9 8 7 6 5 4 3 2 1 0	1 2 3 4 5 6 7 8 9 1	
1	0.74372	·	*****	
2	0.26650	· · ·	****	
3	0.01396	· · · ·	.	
4	0.04114	· · ·	* .	
5	0.09768	· · ·	** .	
6	-0.00031	· · · ·	.	
7	0.13162	· · · ·	*** .	
8	0.24589	· · ·	****	
9	-0.27515	*****	.	
10	-0.04337	. *	.	
11	-0.10618	. **	.	
12	0.13280	· · · ·	***•	
13	-0.04977	. *	.	
14	-0.02398	i . i	•	

Graph 2: Partial Autocorrelation Function for Frequency of Civil Wars

Figure 2 portrays the plot of the differenced version of the "civil wars" variable. The process of differencing evidently helped remove most of the variation and rendered the variable evidently more stochastic than its original version. Thus, this differenced variable contains no time trend; it does, however, possess an intercept, aspect to keep in mind for the ADF. Deriving

the ACF, PACF, and ADF unit root test for the differenced dependent variable will shed light on the nature of the differenced version of the dependent variable.



Figure 2: Frequency of Civil Wars, 1946-2001 – After One Difference

Graph 3 shows the pattern of the ACF. With the differenced version of the variable, the ACF becomes white noise for the most part and only has significance for the first lag and the eighth lag. This indicates that differencing succeeded in removing the heteroskedastic components of the original version of the series.

Graph 3: First-Differenced Autocorrelation Fi	unction for Frequency	y of Civil Conflicts.
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Lag	Covariance	Correlation	-19876543210	0 1 2 3 4 5 6 7 8	9 1 Std Error
1	-31.598732	37034	*****	.	0.134840
2	0.697412	0.00817	.	.	0.152214
3	0.015705	0.00018	· ·		0.152222
4	-1.792944	02101	· ·		0.152222
5	4.160060	0.04876	.	<b> </b> * .	0.152275
6	-11.178176	13101	. ***	.	0.152558
7	-13.154759	15417	. ***		0.154590
8	28.508328	0.33412	· ·	*****	0.157361
9	-22.886437	26823	. *****		0.169770
10	5.332352	0.06250	j .	* .	0.177308
11	-13.370512	15670	. ***	.	0.177708
12	13.756376	0.16122	.	*** .	0.180203
13	-5.456571	06395	i . *	İ .	0.182807

The PACF along with the plot of the variable indicate that one needs to conduct a unit root test with an intercept and time trend and one augmentation term (see Graph 4 for the PACF). The ADF with drift one and (there is an intercept but apparently no time trend based figure 2) with one augmentation meets an appropriate level of statistical significance (0.001) which seems to suggests that the appropriate variable for the proposed model necessitate one difference. Based all the above diagnoses, it appears that the main variable under investigation in this study needs to be differenced one in order to become stationary. Therefore, the model utilized below will use the differenced version of this variable.

Lag	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7 8 9 1
1	-0.37034	*****
2	-0.14948	. ***
3	-0.06128	. * .
4	-0.05020	. *
5	0.02586	. * .
6	-0.12372	. ** .
7	-0.30217	***** .
8	0.17270	. *** .
9	-0.13175	. *** .
10	-0.08898	. ** .
11	-0.25537	****
12	-0.00456	
13	-0.15190	. ***

Graph 4: First-Differenced Partial Autocorrelation Function for Frequency of Civil Conflicts.

After investigating the characteristics of the dependent variable and deciding that it needs one difference to become stationary, we undergo the same process with the independent variable of this study, mainly, hegemonic capabilities.

The same order as for the dependent variable to determine the characteristics of the independent variable (here identified as CAP in SAS). Figure 2 shows the graph of the capabilities of the hegemon – the United States – from 1946 to 2001. Similarly to the dependent variable, the independent variable clearly has both a time trend and an intercept; the only difference comes from the direction of the curve, i.e., the frequency of war seems to constantly increase over time while the relative capabilities of the hegemon follow a declining path over the

time period under investigation. This graphical display may indicate that this variable is nonstationary, which leads to further statistical tests in order to decipher whether such is the case.



In a similar fashion to the dependent variable, the ACF gradually and relatively slowly declines to zero (See Graph 5). Again, this may indicate that the variable is either some sort of AR process or that it is non-stationary in its raw form. Yet, just looking at the ACF does not help generate satisfying conclusions about the nature of the independent variable insofar as the decline to zero seems a little too abrupt to ascertain that the variable is non-stationary.

Lag	Covariance	Correlation	-1987654	3 2 1 0 1 2 3 4 5 6 7 8 9 1	Std Error
1	0.0034478	0.90665		. ***************	0.133631
2	0.0032167	0.84589		**********	0.217290
3	0.0030128	0.79226		**********	0.269758
4	0.0028898	0.75993		*********	0.308523
5	0.0027524	0.72380	· ·	********	0.340311
6	0.0024779	0.65161		**********	0.366772
7	0.0022060	0.58011	.	***********	0.386892
8	0.0019120	0.50281	.	********	0.402125
9	0.0016974	0.44636	.	*******	0.413200
10	0.0015100	0.39707	.	******	0.421722
11	0.0013140	0.34555	.	******	0.428346
12	0.0011055	0.29070	.	*****	0.433296
13	0.00094833	0.24938	.	***** .	0.436764
14	0.00077848	0.20472	1.	****	0.439300

Graph 5: Autocorrelation Function for Hegemon's Capabilities

The PACF function is significant for the first lag only (see Graph 4) and loses significance directly thereafter, which indicates that it is an AR(1) process and that the ADF necessitates one augmentation term. The Augmented Dickey Fuller unit roots test with an intercept and time trend and one augmentation term has a statistical significance of 0.93. We therefore fail to reject the null that the series is non-stationary, which entails that it must first be differenced at least once to see whether such a variable will have become stationary

Lag	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7 8 9 1
1	0.90665	. *************************************
2	0.13413	. *** .
3	0.03773	. * .
4	0.11100	. ** .
5	0.00765	
6	-0.21306	.***
7	-0.09660	. ** .
8	-0.11733	. ** .
9	0.00055	
10	0.02731	. * .
11	0.01388	
12	-0.00774	
13	0.07620	. ** .
14	-0.04847	. *

Graph 6: Partial Autocorrelation Function for Hegemon's Capabilities

Once differenced, the plot of the independent variable still has an intercept but it loses its time trend – therefore, when looking at the ADF unit root test, one needs to look at it with a drift

only (see Figure 4 for the plot of the differenced "hegemon's capabilities" variable). However, the plot of the variable still shows a lot of variation since the series appears to vary greatly at the beginning of the time period under study to then vary at a moderate rate and around zero – which means that the series becomes rather stable towards the end of the time under study. With regards to the statistical tests necessary to address this differenced version, the ADF unit root test will be one with only a drift insofar as no time trend appears on the graphical expression of the series.

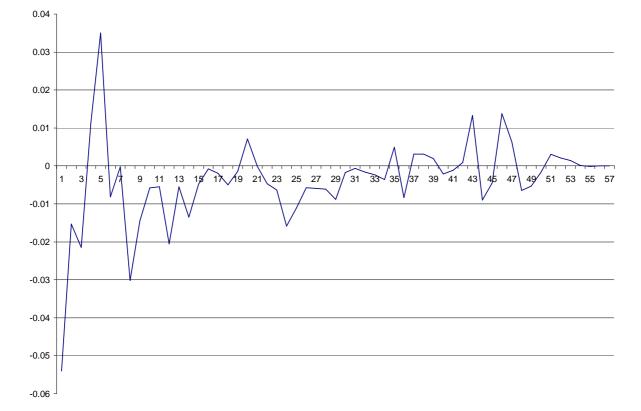


Figure 3: Hegemon's Capabilities, 1946-2001 – After One Difference

After one difference, the ACF for the hegemon's capabilities series becomes mostly white noise though its first lag seems to meet statistical significance (see Graph 7 for the ACF). Consequently, differencing the variable appears to have successfully removed the heteroskedasticity from the original version.

Glabin 7. Thist-Differenced Autocorrelation Function for fregemon's Cabaomities	Graph 7: First-Differenced	Autocorrelation Function	for Hegemon's Capabilities.
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Lag	Covariance	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7 8 9 1	Std Error
1	0.00003156	0.23155	. ****	0.134840
2	9.28911E-6	0.06816	i . i* .	0.141885
3	-0.0000259	19037	. ****	0.142479
4	-0.0000330	24240	.****	0.147031
5	0.00001357	0.09955	. ** .	0.154126
6	0.00001060	0.07781		0.155291
7	0.00001387	0.10176	. ** .	0.155998
8	7.77656E-6	0.05706	. * .	0.157200
9	2.22065E-6	0.01629		0.157576
10	4.57472E-6	0.03357	. * .	0.157607
11	0.00002246	0.16478		0.157737
12	9.51935E-6	0.06985	j . j* .	0.160836
13	9.77856E-6	0.07175	. * .	0.161387
			· ·	-

The PACF shows that the series needs one augmentation for the ADF unit root test (see Graph 8 for the PACF). This test meets statistical significance at the .001 level, leading to the rejection the null that the series is non stationary. In essence, this series needs be differenced once in order to be stationary. In the section that follows – the analysis – the models utilize the first difference version of both the dependent and independent variables.

Graph 8: First-Differenced Partial Autocorrelation Function for Hegemon's Capabilities.

Lag	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1	234567891	
1	0.23155	. ***	**	
2	0.01537	· · ·	.	
3	-0.22139	.****	.	
4	-0.16845	. ***	.	
5	0.24327	. ***	**	
6	0.00473		.	
7	-0.04366	. *	.	
8	0.05534	.  *	.	
9	0.10164	. **	.	
10	0.00460		.	
11	0.18299		*.	
12	0.01631	· · ·	.	
13	0.03122	i . i*	· İ	

#### ANALYSIS

After looking at the characteristics of the dependent and independent variables, the next step requires analyzing whether the two variables fit a transfer function. The computation of the cross-correlation between the frequency of civil wars and the share of total world capabilities possessed by the United States accomplishes this task. Additionally, this study needs to include the lags of the dependent variable that should have an impact on its current values - the number of necessary lagged values emanates from the results of the ACF and PACF for the differenced dependent variable. Referring back to the ACF and PACF of the differenced version of the dependent variable, (Graph 3 and Graph 4), only the first lag is significant in both function and then the series becomes non-significant. These findings does not help identify whether, on its own, the civil war variable necessitate and AR(1), MA(1), or an ARMA process. An information criteria tool, through the use of a MINIC procedure - a procedure that automatically identifies the best fitting model for a stationary variable - assist us in determining the best model (see appendix). The MINIC procedure indicates that the series is an AR(1), therefore, in the event the variables are amenable to a transfer function, the equation will also include the first lag of the dependent variable on the right hand side of the equation. The cross correlation function should help us decide whether 1) the variables fit a transfer function and, if this is the case, 2) how many lags of the independent variable must be included in the model.

Graph 9 shows the results of the cross-correlations between frequency of wars and hegemonic capabilities. Though the cross-correlations between the frequency of wars and the capabilities of the US are all insignificant for the negative lags (which is what we would look for), they are *also* insignificant for positive lags, which seems to indicate that there is no relationship between the two variables.

Lag	Covariance	Correlation	-198765	43210	912	34	56	78	9 :
-13	-0.0072749	06746		. *					
-12	-0.0051275	04755	i	. *	i				
-11	0.0031552	0.02926	i		*				
-10	-0.0026967	02501	i	. *	İ				
-9	-0.0057553	05337	i	. *	i				
-8	-0.0055020	05102	İ	. *	İ				
-7	-0.0044435	04121	i	. *	i				
-6	0.0016754	0.01554	İ		İ				
-5	-0.0045185	04190	İ	. *	İ				
-4	-0.013816	12813	İ	. ***	İ				
-3	0.011691	0.10842	İ		**				
-2	0.021196	0.19656	ĺ		****				
-1	0.0057528	0.05335	İ		*				
0	0.0079291	0.07353	İ		*				
1	-0.016972	15739	Í	. ***	İ				
2	0.0062852	0.05829	ĺ		*				
3	-0.023124	21444	Í	.****	İ				
4	-0.0080739	07487	ĺ	. *	ĺ				
5	0.0081079	0.07519	ĺ		**				
6	0.011253	0.10436	ĺ		**				
7	0.0097008	0.08996		•	**				
8	-0.012533	11623		• **					
9	-0.0067542	06263		• *					
10	0.016511	0.15311		•	***	•			
11	-0.015547	14417		. ***					
12	0.011536	0.10698			**	•			
13	0.0021428	0.01987		•					

Graph 9: Cross-Correlations Between Frequency of War and Hegemonic Capabilities.

In spite of this shortcoming and based on the theory proposed above (mainly that a decrease in the capabilities of the hegemonic power should lead to an increase in the occurrence of civil conflicts), four illustration purposes, this project depicts the results of the AR(1) model as well as that of the transfer function model including the lagged dependent and independent variables. For both models, and contrarily to what the plots seem to suggest, the intercept is irrelevant, therefore, only results excluding the intercept are shown (results with the intercept appear in the appendix). Also, because PRIO has data on conflicts till 2004, I will forecast the next three periods to compare and contrast them with the actual values. The results of the analysis appear in Table 1.

Table 1: Estimates from the AR(1) and Transfer Function)

	AR(1)	Transfer Function
War Lagged	-0.397*** (0.01)	-0.383*** (0.01)

Capabilities Lagged	-	-105.30 (0.20)
2002 Forecost	59.71	59.94
2002 Forecast	(8.65)	(8.69)
2003 Forecast	62.98	
2003 Forecast	(10.12)	-
2004 Forecast	61.69	
2004 Torceast	(12.07)	_
Ν	:	55

\* p<.10; \*\* p<.05; \*\*\* p<.01. The p-values are in parentheses for the coefficient estimates; the standard error is in parentheses for the forecasts.

As table 1 illustrates, hegemonic capabilities do not have a significant effect on the number of civil conflicts in the international system. In spite of this lack of significance, the relationship is in the expectation direction since a one unit increase in the hegemon's capabilities leads to a decrease of 105 civil conflicts in the system. On the other hand, and in both models, the lagged value of the dependent variable seems to explain move of the variation in current values therefore. Both coefficients are significant at the .01 level. Thus, a one unit increase in the change of frequency of civil wars in the most recent period leads to a decrease of .397 in the change of frequency in civil war – which means that roughly 40% of the change in the dependent variable is explained by its lagged value.

The forecasts for 2002 indicate that, based on the AR(1) model, one should expect to observe between 42.41 and 77.01 (by subtracting and adding the double of the standard error to the actual forecast). PRIO accounted for 48 conflicts in 2002, a number included in the range of confidence of the forecast presented here. This first forecast can thus be said to be adequately accurate. For 2003, the model predicts a number of wars that falls between 42.74 and 83.22 – with 95% confidence. PRIO accounted for 46 in 2003, again, a number that falls with the 95% confidence interval here though barely. Additionally, the 2003 forecast contains a huge confidence interval; consequently, one can barely say that this forecast is accurate enough for

there exists a strong gap between 42 wars and 83 wars. As for 2003, the forecasts indicate that the world should experience somewhere between 37.55 and 85.83 – with 95% confidence. For the same year, PRIO reported 91 wars, a number that does not even fall within our level of confidence. A shortcoming of forecasting comes from the fact that they become less and less efficient as we move into the future for there are less and less "actual" observations on which to base the later forecasts. Overall, the range of the forecasts (maybe with the exception of the forecast 2002) appear too wide for anyone to be confident as to their accuracy.

## **DISCUSSION & FUTURE RESEARCH**

This research project attempted to test whether there exists a relationship between the capabilities of the world's hegemonic power and the frequencies of civil conflicts therein. More precisely, this project posited that it expected a negative relationship between the two whereby a decrease in the major power's capabilities should diminish her ability to prevent and pre-empt conflict on the international scene, leading to an increase in civil wars. In order to determine whether the postulated hypothesis holds, data was gathered from competent sources and statistical analyses over the characteristics of those data ensued. The results demonstrate that there is no apparent relationship between the two proposed variables and that, instead, past civil wars explain the current amount of civil wars within the system. These contradictory (or rather non-supportive) findings raise several problems.

First, the models utilized remain very limited in their scope and the inclusion of variables. Much more factors than just past civil wars and the hegemon's power should account for the frequency of civil war at the global level. Second, the number of observations reduces the degrees of freedom and potentially precludes our ability to actually outline an existing relationship. Finally, the characteristics of the variables remain unclear. Both appear to be nonstationary in their raw form (and the unit root tests also seem to suggest this) but they may only necessitate partial difference in order to become stationary (as opposed to the full differencing that was conducted throughout this project). Thus, future research on the topic should attempt to address the points raise in 1) adding further variables through precise theorizing, 2) look for alternative sources of data in order to hopefully cover a longer time period, and 3) adequately manage to find the "true" characteristics of all variables included in the model.

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#### APPENDIX

```
The commands to read the data are:
/* These are the commands for final paper */
title 'Anne Etienne';
filename indata 'C:\Documents and Settings\Anne Etienne\Desktop\Classes\Time
Series\paper.csv';
data war;
infile indata delimiter=',';
input date cap war;
format date year4.;
run;
```

I got the following window upon reading the data into SAS:

```
267 /* These are the commands for final paper */
268 title 'Anne Etienne';
269 filename indata 'C:\Documents and Settings\Anne Etienne\Desktop\Classes\Time Series\paper.csv';
270 data war;
271 infile indata delimiter=',';
272 input date cap war;
273 format date year4.;
274 run;
NOTE: The infile INDATA is:
      File Name=C:\Documents and Settings\Anne Etienne\Desktop\Classes\Time Series\paper.csv,
     RECFM=V, LRECL=256
NOTE: 56 records were read from the infile INDATA.
      The minimum record length was 15.
     The maximum record length was 15.
NOTE: The data set WORK.WAR has 56 observations and 3 variables.
NOTE: DATA statement used (Total process time):
      real time
                         0.01 seconds
      cpu time
                         0.01 seconds
```

I used the following command to check for the unit roots and to get the ACF and PACF of the number of civil wars in the international system as well as for these characteristics with regards to the US capabilities:

```
proc arima;
identify var=war stationarity=(ADF=(0,1,2,3));
identify var=war(1) stationarity=(ADF=(0,1,2,3));
run;
identify var=cap stationarity=(ADF=(0,1,2,3));
identify var=cap(1) stationarity=(ADF=(0,1,2,3));
run;
```

SAS generated the following log window after this command: NOTE: PROCEDURE ARIMA used (Total process time):

```
real time 1:51.62
cpu time 0.61 seconds
283 proc arima;
284 identify var=war stationarity=(ADF=(0,1,2,3));
285 identify var=war(1) stationarity=(ADF=(0,1,2,3));
286 run;
287 identify var=cap stationarity=(ADF=(0,1,2,3));
288 identify var=cap(1) stationarity=(ADF=(0,1,2,3));
289 run;
```

I gathered the following output:

```
Name of Variable = war
Mean of Working Series 41.96429
```

			lard Deviatio er of Observa Autocorr		4.27656 56		
Lag	Covariance	Correlation	-19876		a 1 2 3 4 5	67891	Std Error
0	203.820	1.00000	1 2 2 2 7 2	5 4 5 2 1 0	*********		0
1	151.585	0.74372			   *********	!	0.133631
2	137.010	0.67221		•	   *********	!	0.193936
				•	   ********		
3	116.320	0.57070		•			0.231839
4	104.089	0.51069		•	********		0.255698
5	98.554733	0.48354			********	•	0.273306
6	88.252414	0.43299			*******	•	0.288177
7	91.578922	0.44931	.		*******	•	0.299570
8	104.221	0.51134	.		********	.	0.311371
9	79.657958	0.39082	.		******	.	0.326022
10	73.952578	0.36283	.		******	.	0.334283
11	52.511867	0.25764	i.		****	. i	0.341243
12	56.033528	0.27492	i.		****		0.344699
13	46.541158	0.22834	i ·		   ****		0.348593
14	38.872130	0.19072	· ·		   ****	•	0.351254
14	50.072150		' marks two s	tandand on	1	•	0.551254
		·	Inverse Auto				
	Lag	Correlation	-19876	543210	012345	67891	
	1	-0.32888		******	.		
	2	-0.16277	i	. ***	i.	i	
	3	-0.03030	i	. *	i.	i	
	4	0.05549	i		*	i	
	5	-0.05369	i	•	i -	i i	
	6	0.11450		•	· ·  **		
	7	0.08281		•	•  **		
	8	-0.31286		• *****	•		
					•  **		
	9	0.09923		•	** •		
	10	-0.02293		•	•		
	11	0.15413		•	*** •		
	12	-0.11626		• **			
	13	0.01333		•	.		
	14	0.01444			.		
	Lag	Correlation	Partial Auto		ns 012345	67891	
	1	0.74372	1		*********	****	
	2	0.26650	i		****	i	
	3	0.01396	i			i i	
	4	0.04114	i	•	*  *	i i	
	5	0.09768		•	· ·  **		
	6	-0.00031		•			
	7			•	•  ***		
		0.13162		•	*****   ****		
	8	0.24589		• *****		ļ	
	9	-0.27515					
	10	-0.04337	1	• *	ļ •		
	11	-0.10618		• **			
	12	0.13280		•	*** •		
	13	-0.04977		• *	.		
	14	-0.02398			.		
_			orrelation Ch	eck for Wh:	ite Noise		
То	Chi-	Pr >			<b>.</b>		
Lag	Square	DF ChiSq			Autocorrela		
6	123.16	6 <.0001	0.744	0.672		0.511 0.4	
12	184.46	12 <.0001	0.449	0.511	0.391	0.363 0.2	58 0.275
т.,		•	ed Dickey-Ful Pr < Rho	ler Unit Ro			Pr > F
Тур		Lags Rho		Tau _0 00	Pr < Ta		FI / F
zer	o Mean	0 -0.1435	0.6468	-0.09	0.648		
		1 0.4029	0.7764	0.38	0.791		
		2 0.4320	0.7836	0.46	0.809		
	<b>.</b>	3 0.4202	0.7805	0.48	0.814		
Sin	ngle Mean	0 -11.4120	0.0800	-2.40			0.2798
		1 -6.9170	0.2622	-1.95	0.307	5 2.44	0.4564

# 

Tre	end	1 - 2 -	-4.9435 -4.1770 29.9302 21.3715 22.7070 28.6571	0.4245 0.5053 0.0028 0.0310 0.0215 0.0038	-1.42 -1.23 -4.38 -3.19 -2.90 -2.87	8         0.6           8         0.0           9         0.0           9         0.1	571 051 976 711	1.43 1.12 9.60 5.11 4.21 4.15	0.7099 0.7860 0.0010 0.1695 0.3460 0.3578
		Mean of Standar Number	(s) of Dif Working rd Deviati of Observ	on ations liminated b		0 9	1 .854545 .237115 55 1		
Lag 0 1 2 3 4 5 6 7 8 9 10 11 12 13	Covariance 85.324298 -31.598732 0.697412 0.015705 -1.792944 4.160060 -11.178176 -13.154759 28.508328 -22.886437 5.332352 -13.370512 13.756376 -5.456571	 0.  0.  0.  0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0.	00000 37034 00817 00018 02101 04876 13101 15417 33412 26823 06250 15670 16122 06395	-19876                                     	5 4 3 2 1 ******* **** **** **** **** ****	* * * * * * * * * * * * * * * * * * *	· 5 6 7 8 ********		Std Error 0 0.134840 0.152214 0.152222 0.152222 0.152558 0.154590 0.157361 0.169770 0.177308 0.177708 0.180203 0.182807
	Lag 1 2 3 4 5 6 7 8 9 10 11 12 13	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	I	inverse Auto -1 9 8 7 6                                     	ocorrelatio	ons	**	9 1                       	
	Lag 1 2 3 4 5 6 7 8 9 10 11 12 13	- 0 - 0 - 0 - 0 - 0 - 0 - 0 - 0 - 0 - 0	lation 37034 14948 06128 05020 02586 12372 30217 17270 13175 08898 25537 00456 15190	Partial Auto -19876                                     	5 4 3 2 1 ******** ******* ******* ****** ******	0 1 2 3 4 (		9 1                   	
To Lag 6 12	Chi- Square 9.24 27.06	DF 6 12	Autocor Pr > ChiSq 0.1605 0.0076	CH 		Autocorre 0.000 -0.268		0.049 -0.157	

		Augmented	Dickey-Fuller	• Unit Root	Tests		
Туре	Lags	Rho	Pr < Rho	Tau	Pr < Tau	F	Pr ≻ F
Zero Mean	0	-75.5019	<.0001	-10.54	<.0001		
	1	-90.6174	<.0001	-6.57	<.0001		
	2	-113.219	0.0001	-5.08	<.0001		
	3	-200.656	0.0001	-4.34	<.0001		
Single Mean	0	-75.8928	0.0005	-10.58	0.0001	56.07	0.0010
	1	-94.6124	0.0005	-6.62	0.0001	21.94	0.0010
	2	-126.193	0.0001	-5.13	0.0002	13.18	0.0010
	3	-272.666	0.0001	-4.40	0.0009	9.70	0.0010
Trend	0	-75.9887	<.0001	-10.48	<.0001	54.99	0.0010
	1	-94.5592	<.0001	-6.55	<.0001	21.51	0.0010
	2	-125.227	0.0001	-5.07	0.0007	12.96	0.0010
	3	-252.877	0.0001	-4.33	0.0060	9.60	0.0010

#### Name of Variable = cap Mean of Working Series 0.191778 Standard Deviation 0.061666 Number of Observations 56 Autocorrelations

			Aucocorreit		
Lag	Covariance	Correlation	-1987654	4 3 2 1 0 1 2 3 4 5 6 7 8 9 1	Std Error
Ø	0.0038027	1.00000		***********	0
1	0.0034478	0.90665	ĺ	. ***************	0.133631
2	0.0032167	0.84589	· ·	*********	0.217290
3	0.0030128	0.79226	· ·	*********	0.269758
4	0.0028898	0.75993	.	*******	0.308523
5	0.0027524	0.72380	· ·	*******	0.340311
6	0.0024779	0.65161	.	***********	0.366772
7	0.0022060	0.58011	· ·	*********	0.386892
8	0.0019120	0.50281	.	********	0.402125
9	0.0016974	0.44636	.	*******	0.413200
10	0.0015100	0.39707	1.	******	0.421722
11	0.0013140	0.34555	.	******	0.428346
12	0.0011055	0.29070	.	*****	0.433296
13	0.00094833	0.24938	.	*****	0.436764
14	0.00077848	0.20472	.	****	0.439300
			manke two etas	adand annanc	

#### "." marks two standard errors

#### Inverse Autocorrelations

I

		Inverse	Αl	ıτc	co	orr	٠e.	ιаτ		on	IS									
Lag	Correlation	-198	7	6	5	4	3	2	1	6	) 1	2	3	4	5	6	7	8	9	1
1	-0.40302					**	**	***	**>	*			•							
2	-0.07368								,	*			•							
3	0.07962	1									**		•							
4	-0.02360												•							
5	-0.17478	1						. *	**>	۴İ			•							
6	0.08423	1									**		•							
7	-0.07981								*'	*			•							
8	0.08526	1									**		•							
9	0.03090							•			*		•							
10	-0.02128							•					•							
11	-0.02318												•							
12	0.08148							•			**		•							
13	-0.08823							•	*י	*			•							
14	0.02878										*		•							
		Partial	Δı	ut o	h	orr	م	lat	tid	on	IS									

		Partial	Auto	ocori	rel	lat	:io	ns									
Lag	Correlation	-198	76	54	3	2	1	0 1	2	3	4	5	6	7	8	9	1
1	0.90665							**	**	**:	**:	**>	**>	**:	**:	**	
2	0.13413							**	*	•							
3	0.03773							*		•							
4	0.11100	1						**		•							
5	0.00765	1						1		•							
6	-0.21306	1				**	***			•							
7	-0.09660						**			•							
8	-0.11733	1					**			•							
9	0.00055	1						1		•							
10	0.02731							*		•							
11	0.01388	1						1		•							
12	-0.00774	1						1		•							

	13 14		0.07620 0.04847		•   • *	** .		
τ.				relation Ch	eck for Whi	te Noise		
То	Chi-	55	Pr >					
Lag 6	Square 227.06	DF 6	ChiSq <.0001	0.907		utocorrelat 0.792 0	0.760 0.724	0 652
12	306.24	12	<.0001	0.580	0.846 0.503		0.780 0.724 0.397 0.346	
12	500.24	12	1.0001	0.500	0.303	0.440 0	0.540	0.291
			•	l Dickey-Ful				
Тур		Lags	Rho	Pr < Rho	Tau	Pr < Tau		Pr > F
Zer	ro Mean	0	-1.4471	0.3993	-3.52	0.0007		
		1	-0.9356	0.4804	-2.03	0.0416		
		2 3	-0.9193	0.4832	-1.92 -2.55	0.0534		
Sir	ngle Mean	0	-0.8687 -4.6742	0.4923 0.4525	-2.55	0.0116		0.0010
211	igre mean	1	-2.6664	0.6891	-1.82	0.3674		0.3476
		2	-2.7207	0.6820	-1.82	0.3779		0.4007
		3	-2.1399	0.7550	-1.99	0.2917		0.0870
Tre	nd	0	-6.3891	0.6939	-2.34	0.4036		0.0377
		1	-3.3169	0.9176	-1.08	0.9229		0.8466
		2	-3.4383	0.9113	-1.06	0.9259		0.8546
		3	-0.6222	0.9907	-0.29	0.9888		0.7347
		Ponio	N d(s) of Dif	lame of Vari	able = cap		1	
			of Working			-0	.0039	
			lard Deviati				L1674	
			er of Observ			0.01	55	
				liminated b	v differend	ing	1	
				Autocorr			_	
Lag	Covariance	Corr	elation	-19876	543210	12345	67891	Std Error
õ	0.00013628		1.00000	1		********	******	0
1	0.00003156		0.23155	1	.	****	1	0.134840
2	9.28911E-6		0.06816		.	* •		0.141885
3	-0.0000259		19037		• ****	•		0.142479
4	-0.0000330		24240		•****	•		0.147031
5	0.00001357		0.09955		•	** •		0.154126
6	0.00001060		0.07781		•	** •		0.155291
7	0.00001387		0.10176		•	** • *		0.155998
8	7.77656E-6		0.05706		•	•		0.157200
9 10	2.22065E-6		0.01629		•	•		0.157576 0.157607
10	4.57472E-6 0.00002246		0.03357 0.16478		•	***		
12	9.51935E-6		0.06985		•	*		0.157737 0.160836
13	9.77856E-6		0.07175		•	*		0.161387
19	51770502 0			marks two s	، tandard err	ors.	I	0.10130/
			_					
		Com		inverse Auto			6 7 8 6 1	
	Lag		elation 0.18879	-19876	543216 ****	12345	L ۲ ۵ ۷ ۵	
	1		0.18879		• ****	•		
	3		0.19259		• • • •	• ****		
	4		0.23489	1	•	• ****	1	
	5		0.25171		• *****		ł	
	6		0.04172	i	.	*	İ	
	7		0.01788	i	.		ĺ	
	8		0.06580	i	*	•	i i	
	9	-	0.06419	İ	. *		i	
	10		0.05110	1	.	* .		
	11	-	0.13171		• ***	•		
	12		0.00169		•	•		
	13	-	0.02421		•	•		
			P	artial Auto	correlation	IS		
	Lag	Corr		-1 9 8 7 6			67891	
	1		0.23155			****	· · · · · ·	
	2		0.01537	i	.		i	
	3		0.22139	1	.****		i	

	4		-0.16845	1	***	Ι.		1	
	5		0.24327	i	•	****		i	
	6		0.00473	l		ί.		i	
	7		-0.04366		. *			i	
	8		0.05534	İ	•	*		i	
	9		0.10164	i		** .		i	
	10		0.00460	ĺ	•	į .		Í	
	11		0.18299	1	•	****.		Í	
	12		0.01631	ĺ		į .		Í	
	13		0.03122			* .		Í	
				relation Ch	eck for Wh	ite Nois	e		
То	Chi-		Pr >						
Lag	Square	DF	ChiSq				elations-		
6	10.19	6	0.1168	0.232		-0.190	-0.242	0.100	
12	13.47	12	0.3356	0.102	0.057	0.016	0.034	0.165	0.070
			Augmonted	Dickey-Ful	lon Unit P	oot Toct	<i>c</i>		
Тур	<b>a</b>	Lags	Rho	Pr < Rho	Tau		Tau	F	Pr > F
	o Mean	0	-38.6077	<.0001	-6.91		0001		
2010	5 Hearr	1	-34.0931	<.0001	-4.69		0001		
		2	-66.1990	<.0001	-5.29		0001		
		3	-65.6449	<.0001	-3.99		0002		
Sin	gle Mean	ø	-41.4473	0.0005	-7.11		0001	25.61	0.0010
	5	1	-39.6393	0.0005	-4.93		0002	12.23	0.0010
		2	-97.9559	0.0005	-5.78		0001	16.83	0.0010
		3	-202.261	0.0001	-4.75	0.	0003	11.30	0.0010
Tre	nd	0	-44.6670	<.0001	-7.24		0001	26.92	0.0010
		1	-46.8576	<.0001	-5.15		0005	13.44	0.0010
		2	-170.929	0.0001	-6.32		0001	20.16	0.0010
		3	261.0757	0.9999	-6.05	<.	0001	18.35	0.0010

The command for the cross-correlation function is: proc arima;

identify var=war(1) crosscor=(cap(1));
run;

The log window reads:

NOTE: PROCEDURE ARIMA used (Total process time): real time 1:03:27.54 cpu time 1.14 seconds 290 proc arima; 291 identify var=war(1) crosscor=(cap(1)); 292 run;

I gathered this output:

gamere	za uns outpu	ι.										
			Name of Variable	= war								
		Period(s) of Di	ifferencing		1							
		Mean of Working	g Series	0.854545								
		Standard Deviat	tion		9.237115							
		Number of Obser	rvations		55							
		Observation(s)	eliminated by di-	fferenc	ing 1							
			Autocorrelat:	ions	0							
Lag	Covariance	Correlation	-19876543	3210	1234567891	Std Error						
ō	85.324298	1.00000		1	******	0						
1	-31.598732	37034	**	*****	·	0.134840						
2	0.697412	0.00817		. İ		0.152214						
3	0.015705	0.00018	ĺ .	. İ	·	0.152222						
4	-1.792944	02101		.	.	0.152222						
5	4.160060	0.04876		.	* .	0.152275						
6	-11.178176	13101		. ***	.	0.152558						
7	-13.154759	15417		. ***	.	0.154590						
8	28.508328	0.33412		.	*****	0.157361						
9	-22.886437	26823	.	*****	.	0.169770						
10	5.332352	0.06250	.	Í	* .	0.177308						
11	-13.370512	15670	.	***	.	0.177708						

12 13	13.756376 -5.456571	0.16122 06395		• *	*** ·   ·		:	0.180203 0.182807
			." marks two	standard er	rors			
	مرا	Correlation		tocorrelation 6 5 4 3 2 1 0		5678	01	
	Lag 1	0.51302	-190/	0 5 4 5 2 1 1	0 I Z 3 4  ******			
	2	0.38573		•	  ******		ł	
	3	0.31355	İ		*****		i	
	4	0.34170	i		******		i	
	5	0.29982	İ		*****		İ	
	6	0.35514		•	******		ļ	
	7	0.28997	ļ	•	*****		ļ	
	8	0.10957		•	**  ****			
	9 10	0.25393 0.21695		•	****  ****			
	10	0.21095		•	· · · · · · · · · · · · · · · · · · ·		ł	
	12	0.07808	ł		· ·		ł	
	13	0.08737	İ	•	**		İ	
			Partial Au	tocorrelatio	ns			
	Lag	Correlation		654321		5678	91	
	1	-0.37034		******	.		I	
	2	-0.14948		• ***			ļ	
	3	-0.06128		• *				
	4	-0.05020		• *	·			
	5 6	0.02586 -0.12372		•	* • 			
	7	-0.30217		*****				
	8	0.17270	i		***		i	
	9	-0.13175	i	. ***	į .		i	
	10	-0.08898		• **	.			
	11	-0.25537		****	.		ļ	
	12	-0.00456		•	•			
	13	-0.15190	I	• ***	•		I	
То	Chi-	Auto Pr >	correlation	Check for Wh	ite Noise			
Lag	Square	DF ChiSq			Autocorre	lations		
6	9.24	6 0.1605		0.008	0.000	-0.021	0.049	-0.131
12	27.06	12 0.0076	-0.154		-0.268	0.062	-0.157	0.161
		Var	iable can ha	s been diffe	renced.			
				of war and c				
		Period(s) of [			- F	1		
		Variance of in	•		0	.000136		
		Number of Obse				55		
		Observation(s)	-		cing	1		
	Lag Cova	riance Corre		rrelations 987654	3210	1 2 3 4 5	67891	
	0		.06746	20,004	. *			
			.04755		. *			
	-11 0.0	031552 0	.02926		.  *	•		
			.02501		• *	•		
			.05337		· *  *	•		
			.05102   .04121		• *	•		
			.04121		• "	•		
			.04190		•	•		
			.12813		***	•		
	-3 0.	011691 0	.10842		.  *	•		
			.19656		•  *	***.		
			.05335		•  *	•		
			.07353		•  * ***	•		
			.15739   .05829		• '`TT   *	•		
			.21444		****	•		
			.07487		. *	•		
			.07519		•  *	* .		

0.011253	0.10436	. ** .	
0.0097008	0.08996	. ** .	
-0.012533	11623	. ** .	
-0.0067542	06263	. *  .	
0.016511	0.15311	. *** .	
-0.015547	14417	. ***	
0.011536	0.10698	. ** .	
0.0021428	0.01987	.   .	
	0.0097008 -0.012533 -0.0067542 0.016511 -0.015547 0.011536	0.0097008         0.08996           -0.012533        11623           -0.0067542        06263           0.016511         0.15311           -0.015547        14417           0.011536         0.10698	0.0097008       0.08996       .       **         -0.012533      11623       .       **         -0.0067542      06263       .       *         0.016511       0.15311       .       ***         -0.015547      14417       .       ***         0.011536       0.10698       .       **

"." marks two standard errors

Here are the commands for the MINIC procedure for war: proc arima; identify var=war(1) minic p=(0:4) q=(0:4); run;

I gather the following output:

ather	the following	g output:			
			Name of Variable = war		
		Period(s) of Di	fferencing	1	
		Mean of Working	Series	0.854545	
		Standard Deviat	ion	9.237115	
		Number of Obser	vations	55	
		Observation(s)	eliminated by differen	cing 1	
			Autocorrelations	-	
Lag	Covariance	Correlation	-1987654321	0 1 2 3 4 5 6 7 8 9 1	Std Error
õ	85.324298	1.00000		******	0
1	-31.598732	37034	*****	i. i	0.134840
2	0.697412	0.00817	· ·	i. i	0.152214
3	0.015705	0.00018	i .	i. i	0.152222
4	-1.792944	02101	i .	i. i	0.152222
5	4.160060	0.04876	i .	i* . i	0.152275
6	-11.178176	13101	. ***	i. i	0.152558
7	-13.154759	15417	. ***	i. i	0.154590
8	28.508328	0.33412	i .	*****	0.157361
9	-22.886437	26823	. ****	i. i	0.169770
10	5.332352	0.06250	i .	i* . i	0.177308
11	-13.370512	15670	. ***	i. i	0.177708
12	13.756376	0.16122	i .	*** .	0.180203
13	-5.456571	06395	. *	i. i	0.182807
		"."	marks two standard er	rors	
			Inverse Autocorrelatio	ns	
	Lag	Correlation	-1987654321	0 1 2 3 4 5 6 7 8 9 1	
	1	0.51302		*****	
	2	0.38573	i .	******	
	3	0.31355	· ·	*****	
	4	0.34170	i .	*****	
	5	0.29982	· ·	*****	
	6	0.35514	· ·	*****	
	7	0.28997	i .	*****	
	8	0.10957	· ·	·** •	
	9	0.25393	· ·	****	
	10	0.21695	· ·	****•	
	11	0.20619	· ·	****•	
	12	0.07808	· ·	** .	
	13	0.08737	į .	·** •	
			Partial Autocorrelation		
	Lag	Correlation	-1 9 8 7 6 5 4 3 2 1	0 1 2 3 4 5 6 7 8 9 1	

	1	-	0.37034		****	**				
	2	-	0.14948		. *	**				
	3	-	0.06128			*				
	4	-	0.05020	1		*			Í	
	5		0.02586	1	•	*				
	6	-	0.12372	1	. **					
	7	-	0.30217	1	*****					
	8		0.17270	1		***	•		Í	
	9	-	0.13175	1	. *	**				
	10	-	0.08898	1		**			Í	
	11	-	0.25537	İ	***	**			Í	
	12	-	0.00456	1	•	Í				
	13	-	0.15190	1	. *	**			I	
			Autoco	rrelation	Check for	White	Nois	e		
То	Chi-		Pr >							
Lag	Square	DF	ChiSq			Auto	corr	elations		
6	9.24	6	0.1605	-0.370	0.008	0.0	00	-0.021	0.049	-0.131
12	27.06	12	0.0076	-0.154	0.334	-0.2	68	0.062	-0.157	0.161
					mation Cri					
		Lags	MA 0	MA 1	MA 2		1A 3	MA 4		
		AR Ø			4.231941	4.302				
					4.23183					
					4.30231					
		AR 3			4.36578					
		AR 4			4.424705	4.490	416	4.470538		
				ries model	• •					
			Minimum	Table Valu	ie: BIC(1,0	) = 4.	1320	86		
The comn	The command for estimating the models are:									
proc ari			0							
-	var=war	(1)	roadaarr	= (app / 1)	\ \ ·					
				-(Cap(1						
000100000	n-1 mot		1.2							

```
ident
estimate p=1 method=ML;
estimate p=1 input=(1$ cap) method=ML;
run;
```

```
The log window reads:
```

NOTE: PROCEDURE ARIMA used (Total process time): real time 8:51.10 cpu time 0.60 seconds cpu time 0.60 seconds 305 proc arima; 306 identify var=war(1) crosscorr=(cap(1)); 307 estimate p=1 method=ML; 308 estimate p=1 input=(1\$ cap) method=ML; 309 run;

I get this output:

U	1		Name of Variable	= war					
		Period(s) of D	ifferencing		1				
		Mean of Workin	g Series	0.85	0.854545				
		Standard Devia	tion	9.23	7115				
		Number of Obse	rvations		55				
		Observation(s)	eliminated by di	fferencing	1				
			Autocorrelat	ions					
Lag	Covariance	Correlation	-1987654	3 2 1 0 1 2 3 4 5	67891	Std Error			
0	85.324298	1.00000		*********	******	0			
1	-31.598732	37034	*	*****		0.134840			
2	0.697412	0.00817		.   .		0.152214			
3	0.015705	0.00018		.   .		0.152222			
4	-1.792944	02101		.   .		0.152222			
5	4.160060	0.04876		.  * .		0.152275			
6	-11.178176	13101		• *** •		0.152558			
7	-13.154759	15417		• *** •		0.154590			
8	28.508328	0.33412		• ******		0.157361			
9	-22.886437	26823	.	****		0.169770			
10	5.332352	0.06250	.	* .		0.177308			

11	-13.370512	15670		. ***	0.177708
12	13.756376	0.16122		. ***	0.180203
13	-5.456571	06395		. *	0.182807
		"."	marks two sta	andard errors	

				Inverse Aut	tocorrelati	ons				
	Lag	Corr	elation	-19876	554321		45678	91		
	1		0.51302		•	******	****			
	2		0.38573			*****	**			
	3		0.31355		•	*****				
	4		0.34170		•	******	*			
	5		0.29982		•	*****				
	6		0.35514			******	*			
	7		0.28997		•	*****				
	8		0.10957		•	** ·				
	9		0.25393		•	****				
	10		0.21695		•	****·				
	11		0.20619		•	****•				
	12		0.07808		•	** ·				
	13		0.08737	I	•	** ·		I		
				Partial Aut	tocorrelati	ons				
	Lag	Corr	elation				45678	91		
	8		0.37034	1	*****			 		
	2		0.14948	i	. **	*		i		
	3		0.06128	i	•	*		i		
	4		0.05020	i		*		i		
	5		0.02586	i		·* .		i		
	6	-	0.12372	i	• *	*		i		
	7	-	0.30217	*****						
	8		0.17270	i		*** .		i		
	9	-	0.13175	i	. **	*		i		
	10	-	0.08898	i	. *	*		i		
	11	-	0.25537	i	****	*		i		
	12	-	0.00456	i		į .		i		
	13	-	0.15190	Ì	• **	*		Í		
			Autoco	rrelation (	heck for h	lhita Noi	-			
То	Chi-		Pr >			mille noi.	se			
Lag	Square	DF	ChiSq			-Autocori	relations			
Lag 6	9.24	6	0.1605	-0.370	0.008	0.000		0.049	-0.131	
12	27.06	12	0.0076	-0.154	0.334	-0.268	0.062	-0.157	0.161	
				ble cap has rrelation o						
		Perio	d(s) of Di			cup	1			
			nce of inp				0.000136			
			r of Obser				55			
				eliminated	by differe	encing	1			
		00361	vacion(3)	CIIMINACEU	by drifere		1			

## 

		Cross	corr	ela	nti	.ons	5														
Lag	Covariance	Correlation	-19	8	7	6 5	54	13	32	2 1	L (	01	2	3	4	5	6	7	8	9	1
-13	-0.0072749	06746	1								*										
-12	-0.0051275	04755	Í								*	ĺ									Ì
-11	0.0031552	0.02926	1									*									Ì
-10	-0.0026967	02501	1								*	l l									Ì
-9	-0.0057553	05337	1								*	l l									Ì
-8	-0.0055020	05102	Í								*	ĺ									Ì
-7	-0.0044435	04121	1								*	ĺ									1
-6	0.0016754	0.01554	Í									ĺ									Ì
-5	-0.0045185	04190	1								*	l l									Ì
-4	-0.013816	12813	1							**	**	l l									Ì
-3	0.011691	0.10842	1									**									1
-2	0.021196	0.19656	1									**	**								
-1	0.0057528	0.05335	1									*									1
0	0.0079291	0.07353	1									*									1
1	-0.016972	15739	1							**	**										
2	0.0062852	0.05829	1									*									1
3	-0.023124	21444	1						•*	***	**										1
4	-0.0080739	07487	1								*										
5	0.0081079	0.07519	1									**									Ì
6	0.011253	0.10436										**		•							
7	0.0097008	0.08996	1									**									
8	-0.012533	11623	1							*	**										1
9	-0.0067542	06263	1								*	ĺ									1
10	0.016511	0.15311										**:	*	•							
11	-0.015547	14417	1							**	**										
12	0.011536	0.10698										**		•							
13	0.0021428	0.01987																			
		"." marks t	wo s	tar	nda	rd	er	r	ors	5											

#### . . . . . . . . . . . .

		Мах	imum Like	lihood Est	imation			
			St	andard		Approx		
Parame	eter	Estimat			t Value		Lag	
MU		0.7896	3 0				-	
							1	
			Error Est					
		AIC			395.7243			
		SBC			399.739			
Number of Residuals 55								
		ſ	orrelatio	ns of Para	meter			
		C						
		Dan	= -		AR1 1			
		ANI	<b>⊥</b> ر	0.050	1.000			
		Autoc	orrelatio	n Check of	Residuals			
Chi-		Pr ≻						
Square	DF	ChiSq			Autocor	relations		
4.77	5	0.4441	-0.047	-0.128	-0.048	-0.017	-0.007	-0.234
14.92	11	0.1860	-0.097	0.268	-0.148	-0.081	-0.163	0.111
16.76		0.4707	0.007	-0.035	-0.038	0.121	0.027	0.067
19.14	23	0.6930	-0.024	-0.012	0.068	0.027	-0.117	0.068
			Model for	vaniable	wan			
				vai tabte		1		
				fononcina		-		
		Pel.100(	S) UT D1T	renencting		T		
			Autoregr	essive Fac	tors			
		Fa	•					
	MU AR1,1 Square 4.77 14.92 16.76	AR1,1 Chi- Square DF 4.77 5 14.92 11 16.76 17	Parameter MU         Estimat 0.7896 -0.4033           AR1,1         -0.4033           Cons Vari Std AIC         Cons Vari Std AIC           Vari Std AIC         AIC           C         Par MU           Autoc         Pr >           Square         DF           ChiSq         0.4441           14.92         11           10.1860         16.76           19.14         23           Estimat Period(	Parameter MU         Estimate 0.78963         O           AR1,1         -0.40334         O           Constant Estim Variance Estim Std Error Estim AIC         Std Error Estim Std Error Estim Std Error Estim AIC           SBC         Number of Rest           Number of Rest         Correlation Estimated           Chi-         Pr >           Square         DF           ChiSq	Standard Parameter MU 0.78963 0.83585 AR1,1 -0.40334 0.13202 Constant Estimate Variance Estimate Std Error Estimate AIC SBC Number of Residuals Correlations of Para Estimates Parameter MU MU 1.000 AR1,1 0.030 Autocorrelation Check of Chi- Pr > Square DF ChiSq 4.77 5 0.4441 -0.047 -0.128 14.92 11 0.1860 -0.097 0.268 16.76 17 0.4707 0.007 -0.025 19.14 23 0.6930 -0.024 -0.012 Model for variable Estimated Mean Period(s) of Differencing Autoregressive Fac	Parameter MU         Estimate 0.78963         Error 0.83585         t Value 0.94           AR1,1         -0.40334         0.13202         -3.06           Constant Estimate         1.108127         Variance Estimate         75.05315           Std Error Estimate         75.05315         Std Error Estimate         8.663322           AIC         395.7243         SBC         399.739           Number of Residuals         55         Std Error Estimates         55           Correlations of Parameter Estimates         Estimates         55           Parameter         MU         1.000         0.030           AR1,1         0.030         1.000         AR1,1           MU         1.000         0.030         AR1,1           MU         1.000         0.030         AR1,1           AR1,1         0.030         1.000         AR1,1           Autocorrelation Check of Residuals         Pr >         Square         Pr            Square         DF         ChiSq	Standard       Approx         Parameter       Estimate       Error       t Value       Pr >  t          MU       0.78963       0.83585       0.94       0.3448         AR1,1       -0.40334       0.13202       -3.06       0.0022         Constant Estimate       1.108127         Variance Estimate       75.05315         Std Error Estimate       8.663322         AIC       395.7243         SBC       399.739         Number of Residuals       55         Correlations of Parameter         Estimates       Parameter         MU       1.000       0.030         AR1,1       0.030       1.000         Autocorrelation Check of Residuals         Chi-       Pr >         Square       DF       ChiSq         4.77       5       0.4441       -0.047       -0.128       -0.048       -0.017         14.92       11       0.1860       -0.097       0.268       -0.148       -0.081         16.76       17       0.4707       0.024       -0.012       0.068       0.027         Model for variable war         Estimated Mean       0.78	Standard       Approx         Parameter       Estimate       Error       t Value       Pr >  t        Lag         MU       0.78963       0.83585       0.94       0.3448       0         AR1,1       -0.40334       0.13202       -3.06       0.0022       1         Constant Estimate       1.108127         Variance Estimate       1.108127       Variance Estimate       75.05315         Std Error Estimate       8.663322       AIC       395.7243         SBC       399.739       Number of Residuals       55         Correlations of Parameter         Estimates       Parameter       MU       AR1,1         MU       1.000       0.030       1.000         AR1,1       0.030       1.000         Autocorrelation Check of Residuals         Chi-       Pr >       Square       DF       ChiSq

		Maximum L	ikelihood E	Estimation			
		Standard		Approx			
Parameter	Estimate	Error	t Value	Pr >  t	Lag	Variable	Shift
MU	0.48336	0.92335	0.52	0.6006	Ō	war	0

AR1,1		.38773		13611	-2.85	0.0044	1 0	war	0
NUM1	-88	16126	89.	41047	-0.99	0.3241	0	сар	1
			Con	stant Est	imate	0.670776			
			Var	iance Est	imate	76.53625			
			Std	Error Es	timate	8.7485			
			AIC			390.561			
			SBC			396.528			
			Num	ber of Re	siduals	54			
			Corre	lations o	f Paramete	∽ Estimate	S		
		Va	riable		war	war	cap		
		Pa	rameter		MU	AR1,1	NUM1		
		wa	ır	MU	1.000	0.017	0.360		
		wa	ır	AR1,1	0.017	1.000	-0.034		
		ca	р	NUM1	0.360	-0.034	1.000		
			Auto	correlati	on Check o	f Residual	S		
То	Chi-		Pr >						
Lag	Square	DF	ChiSq			Autoco	rrelations.		
6	4.74	5	0.4489	-0.041	-0.129	-0.030	0.013	0.011	-0.239
12	14.99	11	0.1832	-0.157	0.234	-0.152	-0.091	-0.170	0.112
18	16.99	17	0.4549	0.014	-0.032	-0.033	0.114	0.034	0.091
24	19.49	23	0.6726	-0.002	0.012	0.072	0.028	-0.126	0.062
				Model fo	r variable	war			
			Estima	ted Inter	cept	0.4833	62		
			Period	(s) of Di	fferencing		1		
				Autoreg	ressive Fac	ctors			
			F	actor 1:	1 + 0.387	73 B**(1)			
				Inpu	t Number 1				
			Input	Variable		с	ар		
			Shift				1		

I use the following commands to run the models without an intercept and to forecast the output: proc arima;

1

-88.1613

Period(s) of Differencing

Overall Regression Factor

```
identify var=war(1) crosscorr=(cap(1));
estimate p=1 method=ML noint;
forecast lead=3;
estimate p=1 input=(1$ cap) method=ML noint;
forecast lead=3;
run;
SAS shows the following log window:
       proc arima;
9
10
      identify var=war(1) crosscorr=(cap(1));
11
     estimate p=1 method=ML noint;
12
      forecast lead=3;
     estimate p=1 input=(1$ cap) method=ML noint;
13
      forecast lead=3;
14
15
      run;
I get this output:
                                                     Name of Variable = war
                                Period(s) of Differencing
                                                                                                       1
                                                                                              0.854545
                                Mean of Working Series
                                Standard Deviation
                                                                                              9.237115
                                Number of Observations
                                                                                                      55
                                Observation(s) eliminated by differencing
                                                                                                       1
                                                          Autocorrelations
     Lag
               Covariance
                                  Correlation
                                                      -1 \hspace{0.25cm} 9 \hspace{0.25cm} 8 \hspace{0.25cm} 7 \hspace{0.25cm} 6 \hspace{0.25cm} 5 \hspace{0.25cm} 4 \hspace{0.25cm} 3 \hspace{0.25cm} 2 \hspace{0.25cm} 1 \hspace{0.25cm} 0 \hspace{0.25cm} 1 \hspace{0.25cm} 2 \hspace{0.25cm} 3 \hspace{0.25cm} 4 \hspace{0.25cm} 5 \hspace{0.25cm} 6 \hspace{0.25cm} 7 \hspace{0.25cm} 8 \hspace{0.25cm} 9 \hspace{0.25cm} 1
                                                                                                                       Std Error
                85.324298
                                       1.00000
                                                                                   ******
                                                      0
        0
```

1	-31.598732	37034	*****		0.134840
2	0.697412	0.00817	• 1	•	0.152214
3	0.015705	0.00018	.	•	0.152222
4	-1.792944	02101	.	•	0.152222
5	4.160060	0.04876	.  *	· .	0.152275
6	-11.178176	13101	. ***		0.152558
7	-13.154759	15417	. ***		0.154590
8	28.508328	0.33412	.  *	*****	0.157361
9	-22.886437	26823	. ****		0.169770
10	5.332352	0.06250	.  *	•	0.177308
11	-13.370512	15670	. ***		0.177708
12	13.756376	0.16122	.  *	•**	0.180203
13	-5.456571	06395	. *	•	0.182807
		"." ma	rks two standard erro	ors	

# Inverse Autocorrelations Lag Correlation -1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7 8 9 1

1	0.51302	. ********
2	0.38573	******
3	0.31355	. *****
4	0.34170	******
5	0.29982	*****
6	0.35514	. ******
7	0.28997	. *****
8	0.10957	. ** .
9	0.25393	. ****
10	0.21695	. ****.
11	0.20619	. ****.
12	0.07808	. ** .
13	0.08737	. ** .

		Partial Autocorrelations
Lag	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7 8 9 1
1	-0.37034	*****
2	-0.14948	. ***  .
3	-0.06128	. *  .
4	-0.05020	. *
5	0.02586	. * .
6	-0.12372	. ** .
7	-0.30217	*****
8	0.17270	. *** .
9	-0.13175	. ***  .
10	-0.08898	· **  •
11	-0.25537	****
12	-0.00456	i . i . i
13	-0.15190	. ***

#### Autocorrelation Check for White Noise

То	Chi-		Pr >						
Lag	Square	DF	ChiSq	Autocorrelations					
6	9.24	6	0.1605	-0.370	0.008	0.000	-0.021	0.049	-0.131
12	27.06	12	0.0076	-0.154	0.334	-0.268	0.062	-0.157	0.161

Correlation of war and cap

Correlation of war and cap	
Period(s) of Differencing	1
Variance of input =	0.000136
Number of Observations	55
Observation(s) eliminated by differencing	1

#### Crosscorrelations

		005	30011010			
Lag	Covariance	Correlation	-198	7654321	01234	567891
-13	-0.0072749	06746		• *	.	
-12	-0.0051275	04755		• *	.	
-11	0.0031552	0.02926	1	•	* .	
-10	-0.0026967	02501		• *	.	
-9	-0.0057553	05337		• *	.	
-8	-0.0055020	05102	1	• *	.	
-7	-0.0044435	04121		• *	.	
-6	0.0016754	0.01554			.	
-5	-0.0045185	04190		• *	.	
-4	-0.013816	12813	1	• ***	.	
-3	0.011691	0.10842			** .	
-2	0.021196	0.19656			****.	
-1	0.0057528	0.05335			* .	
0	0.0079291	0.07353		•	* .	
1	-0.016972	15739		. ***	.	
2	0.0062852	0.05829			* .	
3	-0.023124	21444		.****	.	
4	-0.0080739	07487	1	• *	.	
5	0.0081079	0.07519			** .	
6	0.011253	0.10436			** .	
7	0.0097008	0.08996			** .	
8	-0.012533	11623		• **	.	
9	-0.0067542	06263		• *	.	
10	0.016511	0.15311			*** .	
11	-0.015547	14417		. ***	.	
12	0.011536	0.10698			** .	
13	0.0021428	0.01987			.	
		Maximum Li	kelihood	Estimation		
			Standard		Approx	•
	Parameter	Estimate	Error	t Value	Pr >  t	Lag
	AR1,1	-0.39469	0.13177	-3.00	0.0027	1
		Variance Estima		74.90602		
		Std Error Est				
		AIC	Jermace	394.6363		
		SBC		396.6436		
		Number of R	eciduals	550.0450		
			CJIUUUIJ			

Autocorrelation Check of Residuals

То	Chi									
To Lag	Chi- Square	DF	Pr ≻ ChiSq			Autocorr	elations			
6	4.13	5	0.5303	-0.041	-0.109	-0.039	-0.013		-0.225	
12	13.82	11	0.2432	-0.088	0.279	-0.130	-0.054		0.125	
18	15.96	17	0.5264	0.023	-0.015	-0.023	0.132	0.038	0.078	
24	18.31	23	0.7405	-0.016	-0.004	0.073	0.036	-0.109	0.073	
				Model for iod(s) of I mean term	Differenc	ing 1				
			Fa	Autoregree ctor 1: 1						
	0	hc		orecasts fo			onco Limi	.+.		
		bs 57	Forecast 59.7116			95% Confid 42.7485		.6748		
		58	62.9829			43.1541		.8117		
		59	61.6918			38.0304		.3532		
			Max Stan	imum Likel: dard	ihood Est	imation Approx				
Parame	eter E	stimate			Value	Pr >  t	Lag	Variable	Shift	
AR1,1		0.38414	0.1		-2.84	0.0045	1	war	0	
NUM1	-10	5.30265	82.9	4862	-1.27	0.2043	0	cap	1	
			Vari	ance Estima	ate	75.47026				
				Error Esti		8.687362				
			AIC			388.8489				
			SBC			392.8269				
			Numb	er of Resi	duals	54				
			Correl	ations of I	Parameter	Estimates				
			Variabl	e	wa	ir cap				
			Paramet		AR1,					
			war cap	AR1,1 NUM1						
То	Chi			orrelation	Check of	Residuals				
To Lag	Chi- Square	DF	Pr ≻ ChiSq			Autocorr	elations			
6	4.59	5	0.4677	-0.041	-0.126	-0.028	0.017		-0.236	
12	14.57	11	0.2029	-0.156	0.235	-0.146	-0.083		0.120	
18	16.61	17	0.4813	0.021	-0.025	-0.027	0.115	0.037	0.093	
24	19.15	23	0.6926	0.000	0.018	0.078	0.030	-0.124	0.061	
				Model for iod(s) of I mean term	Differenc	ing 1				
			Fa	Autoregres						
			Input V	•	Number 1	сар				
			Shift		-	1				
				s) of Diffe Regression		1 -105.303				
WARNING: Mor The	re values o e value for	•								
	Forecasts for variable war									
	0	bs	Forecast			95% Confid	ence Lim	its		
		57	59.9421			42.9152		.9690		