



Gender/Sex Disparities in the COVID-19 Cascade From Testing to Mortality: An Intersectional Analysis of Swiss Surveillance Data

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Objectives: This study investigates gender and sex disparities in COVID-19 epidemiology in the Canton of Vaud, Switzerland, focusing on the interplay with socioeconomic position (SEP) and age.

Methods: We analyzed COVID-19 surveillance data from March 2020 to June 2021, using an intersectional approach. Negative binomial regression models assessed disparities between women and men, across SEP quintiles and age groups, in testing, positivity, hospitalizations, ICU admissions, and mortality (Incidence Rate Ratios [IRR], with 95% Confidence Intervals [CI]).

Results: Women had higher testing and positivity rates than men, while men experienced more hospitalizations, ICU admissions, and deaths. The higher positivity in women under 50 was mitigated when accounting for their higher testing rates. Within SEP quintiles, gender/sex differences in testing and positivity were not significant. In the lowest quintile, women's mortality risk was 68% lower (Q1: IRR 0.32, CI 0.20–0.52), with decreasing disparities with increasing SEP quintiles (Q5: IRR 0.66, CI 0.41–1.06).

Conclusion: Our findings underscore the complex epidemiological patterns of COVID-19, shaped by the interactions of gender/sex, SEP, and age, highlighting the need for intersectional perspectives in both epidemiological research and public health strategy development.

Keywords: COVID-19 epidemiology, gender and sex, social determinants of health, intersectionality, public health

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Received: 08 January 2024

Accepted: 10 May 2024

Published: 21 May 2024

Citation:

Auderset D, Amiguet M, Clair C, Riou J, Pittet V, Schwarz J and Mueller Y (2024) Gender/Sex Disparities in the COVID-19 Cascade From Testing to Mortality: An Intersectional Analysis of Swiss Surveillance Data. *Int J Public Health* 69:1607063. doi: 10.3389/ijph.2024.1607063

INTRODUCTION

The COVID-19 pandemic has had heterogeneous impacts, with certain populations being disproportionately affected. The literature on COVID-19 indicates that socioeconomically disadvantaged groups face higher risks of contracting the virus and experiencing severe outcomes such as hospitalization and mortality [1–5]. This risk is linked to socioeconomic determinants, where limited income and education create conditions that elevate exposure risk and susceptibility to infection [1]. In many countries, studies have highlighted socioeconomic

disparities in the COVID-19 cascade, both nationally [5–8] and regionally [4, 9], indicating that neighborhood-level socioeconomic vulnerability shapes these disparities [10]. Additionally, substantial gender differences, particularly in labor and family domains, impact individuals throughout their lifespan, contributing to gendered socioeconomic inequalities that affect health [11].

Globally, men were more likely to develop severe forms of COVID-19, resulting in higher hospitalizations and mortality rates compared to women [12–18]. The origins and pathways of these disparities are rooted in a complex interplay of gender-specific social processes and sex-related biological attributes. Gender, as a major social determinant of health, shapes life experiences and health outcomes through systematic differences in roles, responsibilities, access to power, and opportunities between women and men [19, 20]. Gender inequalities in health arise from the intricate interaction of multiple factors, including differential exposure to health risks, health-related behaviors, access to healthcare, and gender biases in healthcare and research [19, 21]. Exploring gender and sex reveals nuanced pathways impacting the COVID-19 progression from testing rates to mortality. Gender affects individuals' health behaviors, access to healthcare, occupational exposures, and adherence to public health measures, thereby potentially affecting COVID-19 exposure, testing rates, positivity, and the burden of disease [15, 22–25]. Meanwhile, sex-related differences, such as immune responses and hormone levels, primarily influence susceptibility, severity, and mortality rates of COVID-19 [13, 26, 27]. Studies suggest that hormones like oestrogens and progesterone, typically higher in women, offer protection against viral infections, whereas testosterone, predominant in men, may have the opposite effect [28, 29]. Additionally, men often exhibit higher ACE-2 receptor levels, used by SARS-CoV-2 to enter cells, potentially explaining the more severe infection cases [26, 28, 30].

There are significant variations in the influence of gender, sex, and socioeconomic conditions on health throughout the lifespan, shaped by factors including evolving gender norms, fluctuating economic resources, and sex-specific developmental stages such as puberty or menopause. These variations underscore the importance of intersectional approaches to deepen our understanding of COVID-19 epidemiology [28, 31]. Building on Kimberlé Crenshaw's work, the concept of intersectionality emerges as a critical framework that recognizes that women and men do not constitute homogenous groups [32]. It argues that their health should be investigated by simultaneously considering other major determinants of health, such as age and socioeconomic status [19, 33]. An intersectional perspective highlights the intricate ways in which socially constructed categories—rooted in structural power relations—intersect at various levels to produce nuanced layers of advantage or disadvantage [19, 34, 35]. This approach moves beyond the consideration of isolated risk factors [36] and focuses on identifying modifiable causes for health inequalities, offering major insights for formulating equitable public health policies and interventions [37].

This study analyzes surveillance data from the canton of Vaud, located in the southwestern part of Switzerland within the French-speaking region. As one of Switzerland's larger cantons by population and area, it encompasses diverse urban and rural settings, representing approximately one-tenth of the national population. The primary objective is to explore gender and sex disparities in the COVID-19 epidemiology cascade, from testing to mortality, including test positivity, hospitalization, and intensive care unit (ICU) admission. We examined these disparities in the context of key social determinants of health, focusing on neighborhood-based socioeconomic position (SEP) and age, through an intersectional analytical approach. We aim to uncover the complex dynamics underlying these disparities and enhance our understanding of COVID-19's broader epidemiology. Such knowledge, by considering gender and sex influences, vulnerabilities, and diverse social determinants of health is essential for developing targeted and effective public health strategies to address COVID-19 and guide responses to future pandemics [28].

METHODS

Study Design and Setting

This observational retrospective study analyzed COVID-19 surveillance data from March 2020 to the end of June 2021 of the population residing in the canton of Vaud. The first epidemic wave in Switzerland spanned from February to May 2020, characterized by low testing capacities with RT-PCR tests [4, 5]. The tested population primarily included symptomatic individuals, those with known risk factors (e.g., people with comorbidities), and healthcare workers [4]. Testing was expanded on June 24, 2020, to include mildly symptomatic individuals and close contacts of infected individuals, with test costs reimbursed [5]. Vaccinations began in December 2020, initially for vulnerable groups, and expanded with the opening of vaccination on 11 January 2021 [38, 39]. By June 2021, 85% of those aged 75 or older, and 53% of those aged 18 to 49 had received at least one vaccine dose [40].

Data

Within the Swiss federal state, the Federal Office of Public Health (FOPH) oversees the monitoring of transmissible diseases, including COVID-19, in collaboration with cantonal authorities, through mandatory reporting of infectious diseases [38]. Entities authorized to conduct SARS-CoV-2 testing (RT-PCR and rapid antigen tests), such as general practitioners, pharmacies, and testing centers, had to notify each test (negative and positive) to the FOPH. Hospitalizations (lasting at least 24 h) and ICU admissions due to COVID-19 were reported by hospitals. Probable or confirmed COVID-19-related deaths were reported to cantonal health authorities (see **Supplementary Table S1** for case definitions). Population data as of 31 December 2020, were obtained from the cantonal office of statistics [41]. The SEP of notified individuals was determined using the Swiss-SEP, an area-based indicator [42, 43]. Detailed

information on the geocoding procedures is provided in **Supplementary Section S2**.

The study period spanned 69 weeks from 2 March 2020 (first notified cases in Vaud canton), to 27 June 2021, marking the cessation of the cantonal hospital's surveillance system. Due to inconsistent negative test reporting prior to 24 May 2020, the dataset for the total number of tests was limited to the period from 27 May 2020, to 27 June 2021, covering a span of 57 weeks. Notifications included the date, test result (positive, negative), date of birth or age, and residential address. Additionally, for hospitalization, death, and PCR test notifications, administrative sex was recorded, restricted to "female" or "male" in Switzerland. For rapid antigenic tests, however, gender identity was recorded, offering options "other", "women", and "men". Consequently, we have referred to this variable as "gender/sex" throughout our analysis to better acknowledge the inclusion of both aspects in our data aggregation. Duplicated notifications, records with invalid residential addresses, and those missing age or gender/sex information were excluded. Additionally, notifications with "other" as gender/sex (0.001% of total tests) were excluded. Age was grouped into eight categories of 10-year age bands and 80 and above. For hospitalizations, ICU admissions, and deaths, ages 0–59 years were combined due to low number of events in this age range.

The Swiss Socio-Economic Position (Swiss-SEP)

The Swiss-SEP, an area-based socio-economic position index centered on each residential building and incorporating neighborhood information from the surrounding 50 households, was developed by the Swiss National Cohort [43] (detailed in **Supplementary Section S3**). The Swiss-SEP index, derived through principal component analysis, aggregates neighborhood-level data from the 2000 census and 2012–2015 annual micro-census. This index utilizes key indicators as proxies for SEP: median rent per square meter (income proxy), proportion of households led by individuals with a primary education or less (education proxy), proportion of households headed by individuals in manual or unskilled jobs (occupation proxy), and the average number of persons per room (crowding proxy). The index scores range from 0 to 100, with higher values indicating higher SEP [42]. There are 115,596 SEP neighborhoods within the geographical boundaries of Vaud canton.

Residential coordinates of each notification were matched with the nearest SEP neighborhood, and SEP index values were categorized into quintiles from one (lowest) to five (highest). Non-residential addresses, such as schools or nursing homes, and addresses with only ZIP code information, were assigned the average SEP of their ZIP code area. Regarding total and positive test notifications, 94% and 92%, respectively, were successfully geocoded, thus assigned an address-based SEP (See **Supplementary Table S4**). However, notifications for hospitalization and ICU admission, that contained only ZIP code information, did not receive a SEP assignment due to the method of assigning average SEP scores to ZIP code areas, which

tends to centralize distribution around mean values, thereby reducing variability at the extremes of the SEP quintiles. Death notifications that could not be geocoded (39%) were likewise excluded from analyses requiring SEP attribution, due to similar concerns regarding the accuracy of SEP assignment.

Statistical Analysis

The distribution of notifications stratified by gender/sex, across age groups, and SEP quintiles were described. Incidence rates of tests, positive tests, hospitalizations, ICU admissions, and deaths were calculated weekly per 100,000 persons, stratified by gender/sex categories. Cumulative incidence rates over the study period were similarly computed. Negative binomial regression models were used to examine the incidence rate ratios (IRR) with 95% confidence intervals (CIs) between women and men, with interaction terms between gender/sex and age groups and between SEP quintiles. These models, which can handle overdispersion of residuals, included denominators as offsets, with corresponding age and sex structure of the general population as of 31 December 2020, serving as the base for all outcomes. Specifically for positive tests, an additional negative binomial model using the total number of tests as the denominator was formulated to investigate gender/sex-specific test positivity ratios. A similar methodology was applied for ICU admissions, with hospitalizations serving as the offset. Sensitivity analyses were conducted on death notifications, incorporating notifications from institutional locations, followed by a comprehensive analysis of all death notifications, including those not precisely geocoded. Statistical analyses were conducted using R statistical software [44], and negative binomial models estimated using the MASS package [45].

Data Reporting Standards

This research aligns with the Sex and Gender Equity in Research (SAGER) guidelines, which advocate for systematic integration of sex and gender considerations into research design, analysis, and reporting [46]. Consequently, we will discuss the sex and/or gender-related mechanisms potentially influencing the findings reported within the context of Switzerland. In this paper, the term "gender/sex" is used to acknowledge the complex interplay between these concepts from a theoretical perspective [47]. This terminology effectively reflects the varied nature of surveillance data analyzed, where the indicator might represent either administrative sex or gender identity, depending on the notification type. From a methodological standpoint, gender/sex acts as a proxy capturing both gender-related aspects (e.g., behaviors) and sex-related biological factors (e.g., hormonal differences), which may impact the outcomes studied.

RESULTS

By the end of 2020, Vaud population was 815,300, comprising 412,599 women (50.6%) and 402,701 men (49.4%) (**Table 1**). From March 2020 to June 2021, a total of 885,925 SARS-CoV-2 tests, 96,963 positive tests, 6,356 hospitalizations, 1,134 ICU admissions and 1,175 deaths (before excluding non-geocoded

TABLE 1 | Distribution of age and socioeconomic position (SEP), stratified by gender/sex categories, Canton of Vaud surveillance data, from 2nd March 2020 to 27 June 2021, Switzerland.

Age groups	Population (N = 815,300)		Total tests* (N = 885,925)		Positive tests (N = 96,963)		Hospitalisations (N = 6,356)		ICU admissions (N = 1,134)		Deaths (N = 1,175)	
	Women N = 412,599	Men N = 402,701	Women N = 463,105	Men N = 422,820	Women N = 50,296	Men N = 46,667	Women N = 2,720	Men N = 3,636	Women N = 334	Men N = 800	Women N = 558	Men N = 617
0-9	42,502 (10%)	44,319 (11%)	10,675 (2.3%)	12,225 (2.9%)	620 (1.2%)	649 (1.4%)	30 (1.1%)	31 (0.9%)	0 (0%)	1 (0.1%)	0 (0%)	0 (0%)
10-19	44,291 (11%)	47,101 (12%)	47,039 (10%)	47,542 (11%)	4,644 (9.2%)	4,743 (10%)	50 (1.8%)	48 (1.3%)	1 (0.3%)	1 (0.1%)	0 (0%)	0 (0%)
20-29	52,178 (13%)	54,565 (14%)	84,209 (18%)	76,947 (18%)	8,854 (18%)	7,987 (17%)	128 (4.7%)	69 (1.9%)	8 (2.4%)	5 (0.6%)	0 (0%)	1 (0.3%)
30-39	59,329 (14%)	59,317 (15%)	94,883 (20%)	81,253 (19%)	8,799 (17%)	7,896 (17%)	173 (6.4%)	121 (3.3%)	13 (3.9%)	12 (1.5%)	0 (0%)	0 (0%)
40-49	59,869 (14%)	58,157 (15%)	74,146 (16%)	67,652 (16%)	8,410 (17%)	7,387 (16%)	160 (5.9%)	250 (6.9%)	15 (4.5%)	52 (6.5%)	1 (0.3%)	1 (0.3%)
50-59	57,692 (14%)	56,987 (14%)	62,506 (13%)	59,343 (14%)	7,692 (15%)	7,494 (16%)	299 (11%)	541 (15%)	60 (18%)	147 (18%)	3 (0.9%)	5 (1.3%)
60-69	40,758 (9.8%)	38,247 (9.6%)	36,464 (7.9%)	37,232 (8.8%)	4,260 (8.5%)	4,734 (10%)	384 (14%)	740 (20%)	88 (26%)	238 (30%)	13 (4.1%)	27 (6.8%)
70-79	34,053 (8.2%)	27,488 (6.9%)	25,042 (5.4%)	23,742 (5.6%)	2,967 (5.9%)	3,170 (6.8%)	542 (20%)	873 (24%)	93 (28%)	242 (30%)	54 (17%)	103 (26%)
80+	24,246 (5.8%)	14,201 (3.5%)	28,141 (6.1%)	16,884 (4.0%)	4,050 (8.1%)	2,607 (5.6%)	954 (35%)	963 (26%)	56 (17%)	102 (13%)	246 (78%)	259 (65%)
Quintile of SEP												
Q1 (lowest)	83,022 (20%)	80,913 (20%)	79,987 (17%)	74,721 (18%)	10,334 (21%)	9,382 (20%)	NA	NA	NA	NA	50 (16%)	92 (23%)
Q2	82,756 (20%)	80,512 (20%)	88,181 (19%)	80,457 (19%)	10,560 (21%)	9,708 (21%)					78 (24%)	72 (18%)
Q3	82,951 (20%)	80,319 (20%)	95,105 (21%)	83,934 (20%)	10,649 (21%)	9,935 (21%)					69 (22%)	88 (22%)
Q4	83,236 (20%)	79,726 (20%)	101,244 (22%)	92,853 (22%)	9,927 (20%)	9,226 (20%)					56 (18%)	77 (20%)
Q5 (highest)	82,953 (20%)	78,912 (20%)	98,588 (21%)	90,855 (21%)	8,826 (18%)	8,416 (18%)					64 (20%)	67 (17%)

Notes: n (%); SEP, socioeconomic position; NA, Not Applicable. Death's distribution across SEP quintiles only includes geocoded notifications. * For total tests, period goes from 27 May 2020, to 27 June 2021.

death notifications) were notified and met eligibility criteria (see **Supplementary Figure S5**). Although women had more tests and positive results, the majority of hospitalizations, ICU admissions, and deaths occurred among men.

In the canton of Vaud, 38% of women and 34% of men were aged 50 and above. Among women, this age group accounted for 33% of all tests and 38% of positive tests, but represented 80% of hospitalizations, 89% of ICU admissions, and 99.8% of deaths. Likewise, men aged 50 and older accounted for 32% of tests, 39% of positive tests, 86% of hospitalizations, 91% of ICU admissions, and 99.4% of deaths.

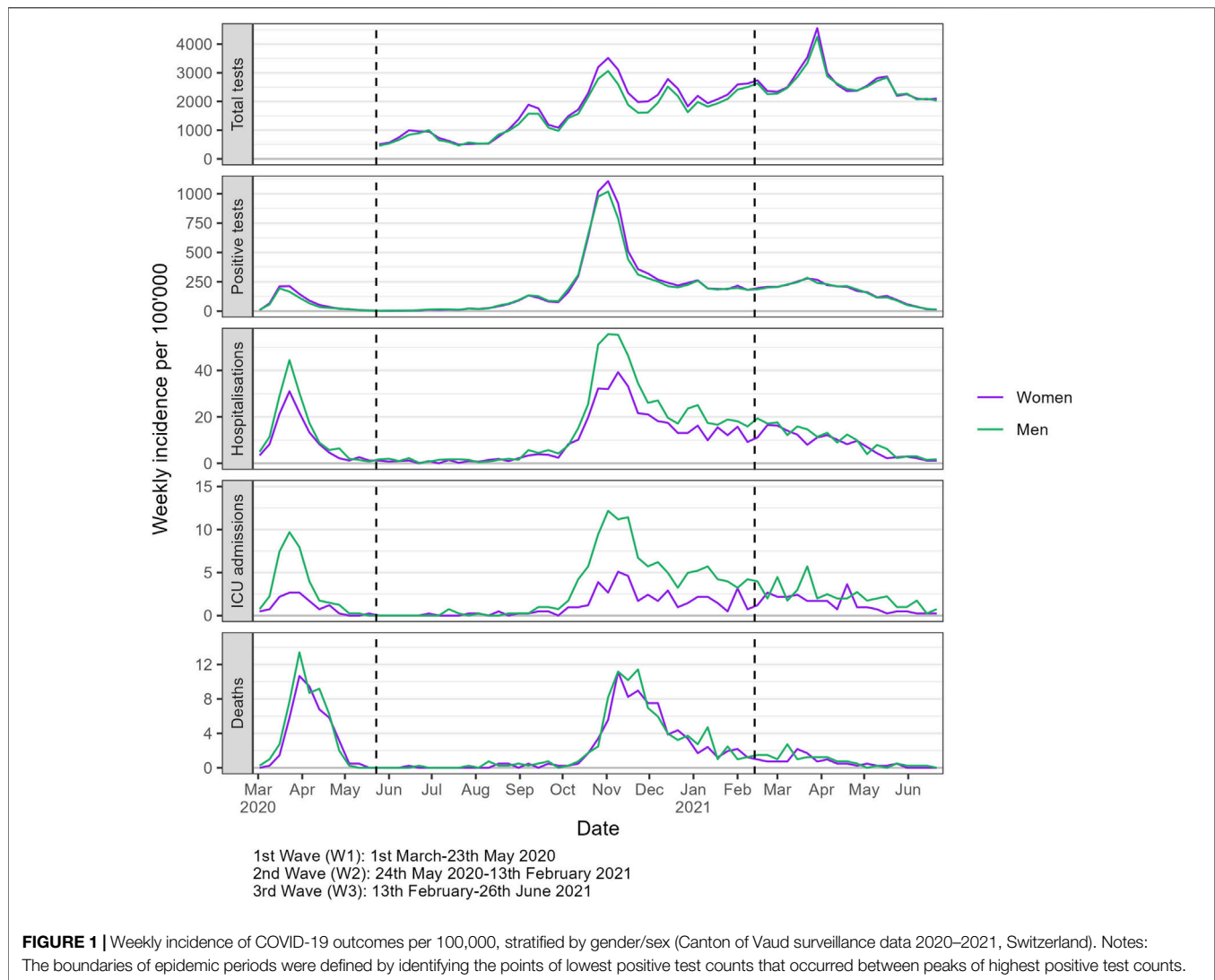
For total tests, 17% of women and 18% of men were in the lowest socioeconomic quintile (Q1), while 21% of tests for both women and men occurred in the highest quintile (Q5). Regarding positive tests, 21% for women and 20% for men were recorded in Q1, with 18% in Q5 for both. In terms of mortality, 18% of men who died were in Q1, and 14% in Q5. Among women, 12% of deaths occurred in Q1 and 15% in Q5.

The weekly incidence of outcomes per 100,000 showed distinct patterns between women and men throughout the study period (**Figure 1**). Women had higher incidence rate of tests and positive tests compared to men, especially during the second wave of the pandemic. In contrast, men had higher incidence rates of hospitalizations and ICU admissions throughout the study period, though the disparity in mortality rates was less pronounced. During the third wave, while testing rates peaked, both severe outcomes and positivity rates were comparatively lower.

The cumulative incidences of outcomes across age groups, gender/sex categories, and SEP quintiles revealed distinct patterns (**Figure 2**). Individuals aged 20 to 39 were the most tested group, whereas children under 10 were the least tested. Testing rates were higher for people aged 80 and above compared to those aged 60-69 and 70-79. Similar patterns emerged across age groups concerning positivity. For severe outcomes, prominent age-related trends were observed, with older age groups experiencing higher incidence rates, though ICU admission were less frequent among those aged 80 and above. Men experienced higher incidence rates of hospitalization, ICU admission, and death than women.

In terms of SEP quintiles (**Figure 2B**), the cumulative incidence of testing progressively increased from Q1 to Q5. For the cumulative incidence of positive tests, the three lowest SEP quintiles (Q1-Q3) showed similar rates, with lower rates observed in the two highest quintiles (Q4-Q5). A consistent trend was observed, with women having higher cumulative incidence of both tests and positive tests across all quintiles, except in Q5 where the positivity incidence was comparable between women and men. For death notifications, men's cumulative mortality rate appeared to decrease from Q1 to Q5, whereas for women, the mortality rate was lowest in Q1 and Q4. Men displayed higher mortality rates across all SEP quintiles, except in Q2.

Regression analyses demonstrated distinct testing patterns between women and men across age groups. Notably, women aged 20-29 (IRR 1.14, CI 1.07-1.22) and 30-39 (IRR 1.16, CI 1.09-1.24) displayed a significantly higher likelihood of

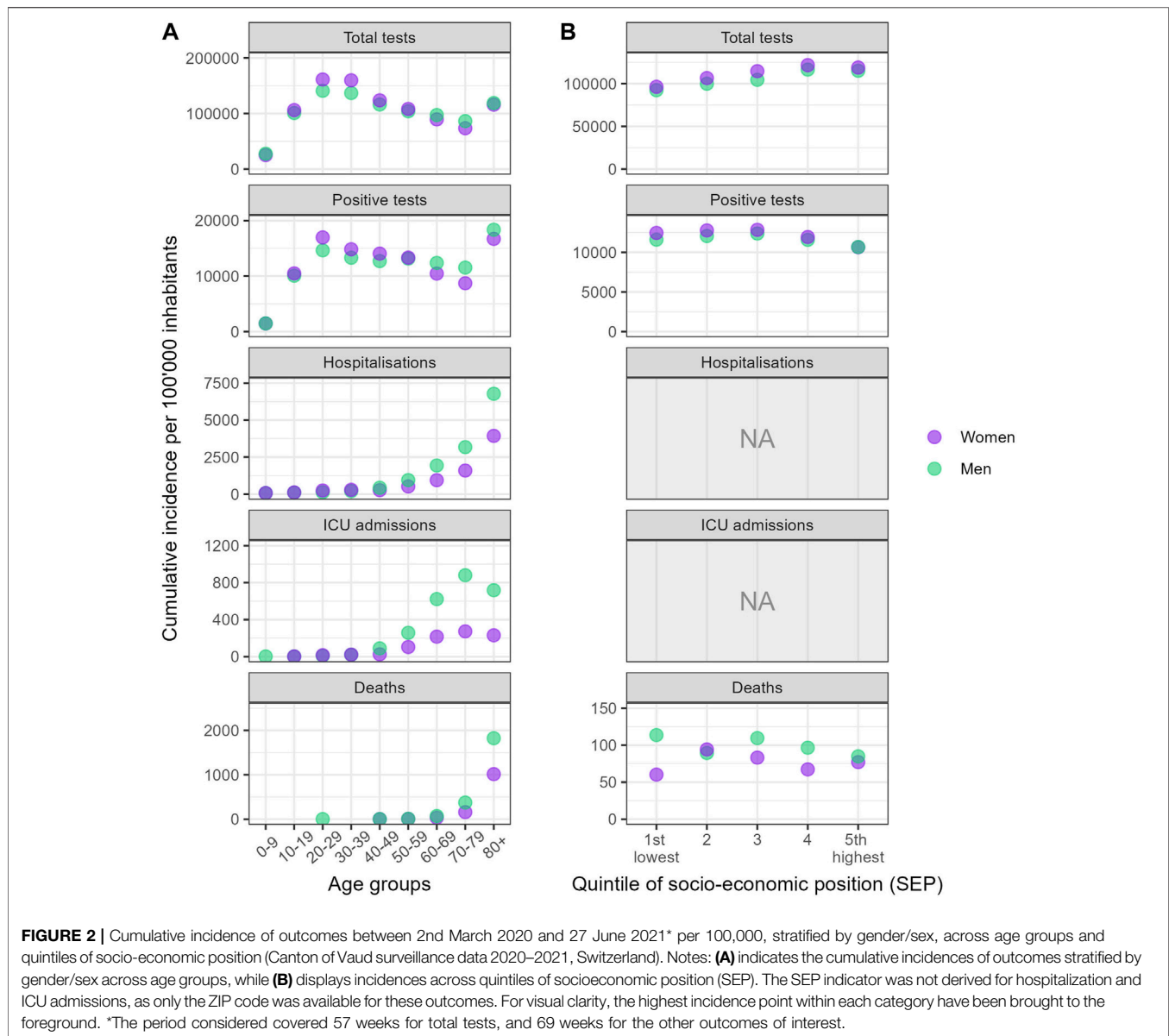


undergoing testing compared to men in the same age groups (Figure 3, left panel). Conversely, girls under 10 (IRR 0.91, CI 0.85–0.97) and women aged 60–69 (IRR 0.92, CI 0.86–0.98) and 70–79 (IRR 0.85, CI 0.80–0.91) were less likely to get tested compared to their male counterparts. For incidence of positive tests per population, similar gender/sex trends were observed across age groups (Figure 3, center panel). However, these differences were no longer significant when adjusting for the initial gender/sex differences in testing, as indicated by the regression results of positive tests per test (Figure 3, right panel). An exception was observed among individuals aged 60 and older, where women were less likely to test positive per test compared to men. Specifically, women had an IRR of 0.92 in the 60–69 age group (CI 0.86–0.98), 0.89 in the 70–79 age group (CI 0.82–0.95), and 0.83 among those aged 80 and older (CI 0.86–1.00), indicating a lower likelihood of a positive result when tested. Moreover, when comparing women and men in similar SEP quintiles (Figure 3, red coefficients), no statistically

significant differences in testing and positive testing rates were found, except for women in Q3 who presented a slightly reduced probability of testing positive per test compared to their male counterparts.

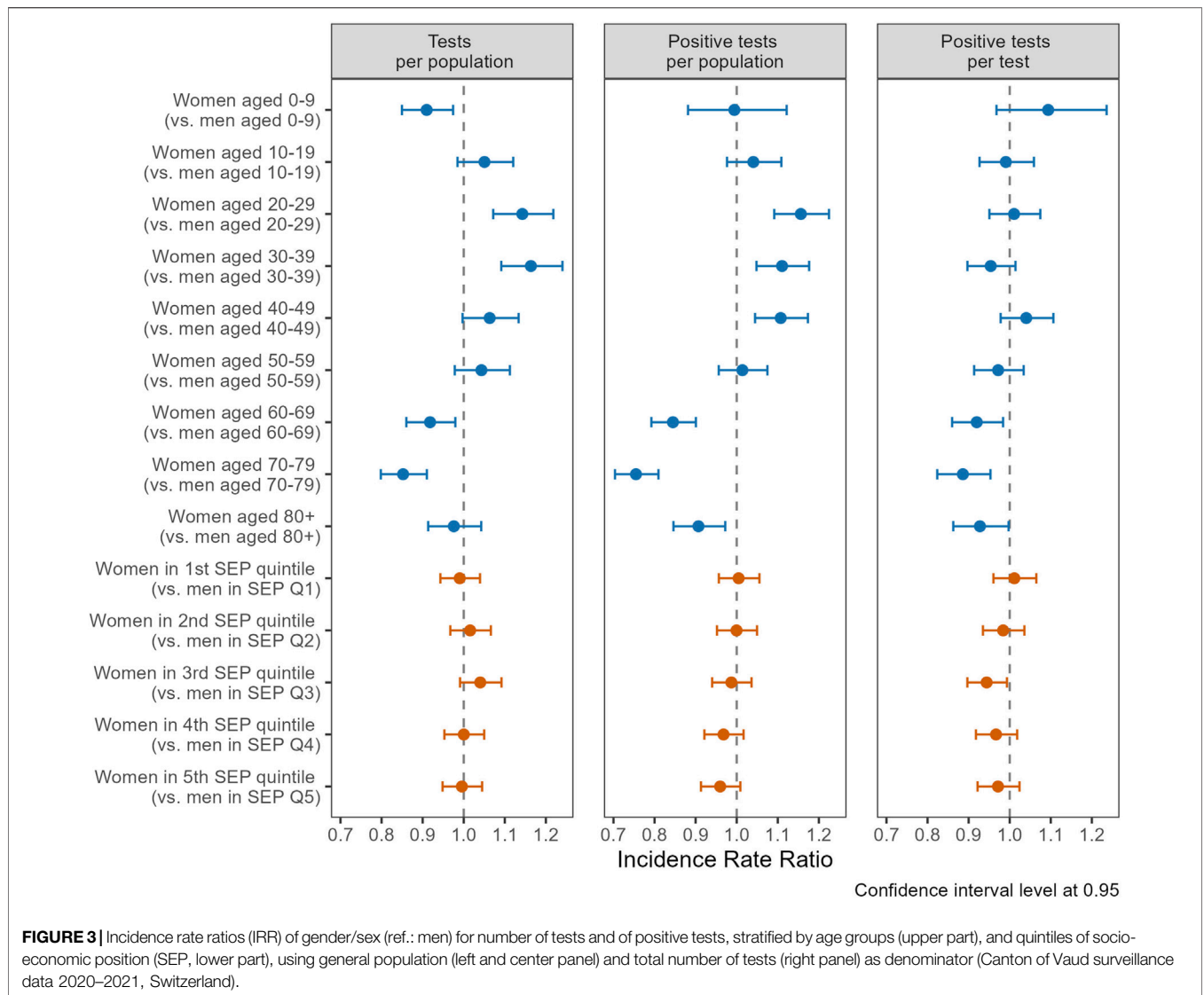
In regression models without an interaction term for gender/sex categories (Supplementary Table S6), individuals in Q5 were notably more likely to undergo testing (IRR 1.25, CI 1.19–1.30) compared to those in Q1. Conversely, individuals in Q5 showed a decreased likelihood of testing positive per person (IRR 0.89, CI 0.85–0.95) and testing positive per test (IRR 0.71, CI 0.68–0.74).

Age was the strongest predictor for hospitalisations, ICU admissions and deaths, increasing age being associated with higher likelihoods of these events, as shown by the regression analysis in models without an interaction term by gender/sex (Supplementary Table S6). Analyzing the interaction of gender/sex with age (Figure 4A), women exhibited lower probabilities of COVID-19 hospitalization than men across all age cohorts,



although the differences were not statistically significant. The IRR for women up to 59 was 0.80 (CI 0.52–1.22), 0.49 (CI 0.18–1.36) for those 60–69, 0.50 (CI 0.18–1.38) for the 70–79 age group, and 0.58 (CI 0.21–1.61) for those 80 and above. Statistical significance was achieved for ICU admissions in the under-60 cohort, where women had a 55% decreased risk (IRR 0.45, CI 0.23–0.86). Women consistently showed a lower risk of ICU admissions compared to men when hospitalized (**Figure 4B**). Women under 60 presented an IRR of 0.59 (CI 0.44–0.78), denoting a 41% lower risk. In the 60–69 age group, the IRR was 0.71 (CI 0.49–1.01), with the risk reduction becoming more pronounced with advancing age. Women aged 70–79 had an IRR of 0.62 (CI 0.43–0.88), while those aged 80 and over had an IRR of 0.56 (CI 0.37–0.85), mirroring the risk reduction observed in the youngest age group.

Regarding mortality, individuals in Q5 had a lower likelihood of death (IRR 0.71, CI 0.54–0.95) compared to those in Q1, as indicated by the regression models without a gender/sex interaction term (**Supplementary Table S6**). This association between SEP and death persisted in the sensitivity analysis that included nursing home residents and remained robust when extended to include non-precisely geocoded death notifications (**Supplementary Material S7**). Lower mortality rates among women were noted across all age groups (**Figure 5**). Women demonstrated a reduced mortality risk compared to men of 55% at ages 60–69 (IRR 0.45, CI 0.23–0.88), 58% at ages 70–79 (IRR 0.42, CI 0.30–0.59), and 45% for those aged 80 and above (IRR 0.55, CI 0.46–0.66). Regarding SEP, the gender/sex disparities in mortality were more pronounced in Q1, with



women having a 68% reduction in mortality risk (IRR 0.32, CI 0.20–0.52); and these disparities were not statistically significant in Q2 and Q5. Exploring the combined effects of gender/sex, age, and SEP on mortality among older groups (70–79 and 80+), a triple interaction term was employed (Supplementary Figure S8). Our findings indicate a reduction in gender/sex mortality disparities with increasing SEP, as the IRR tends toward unity from the lowest to highest quintiles.

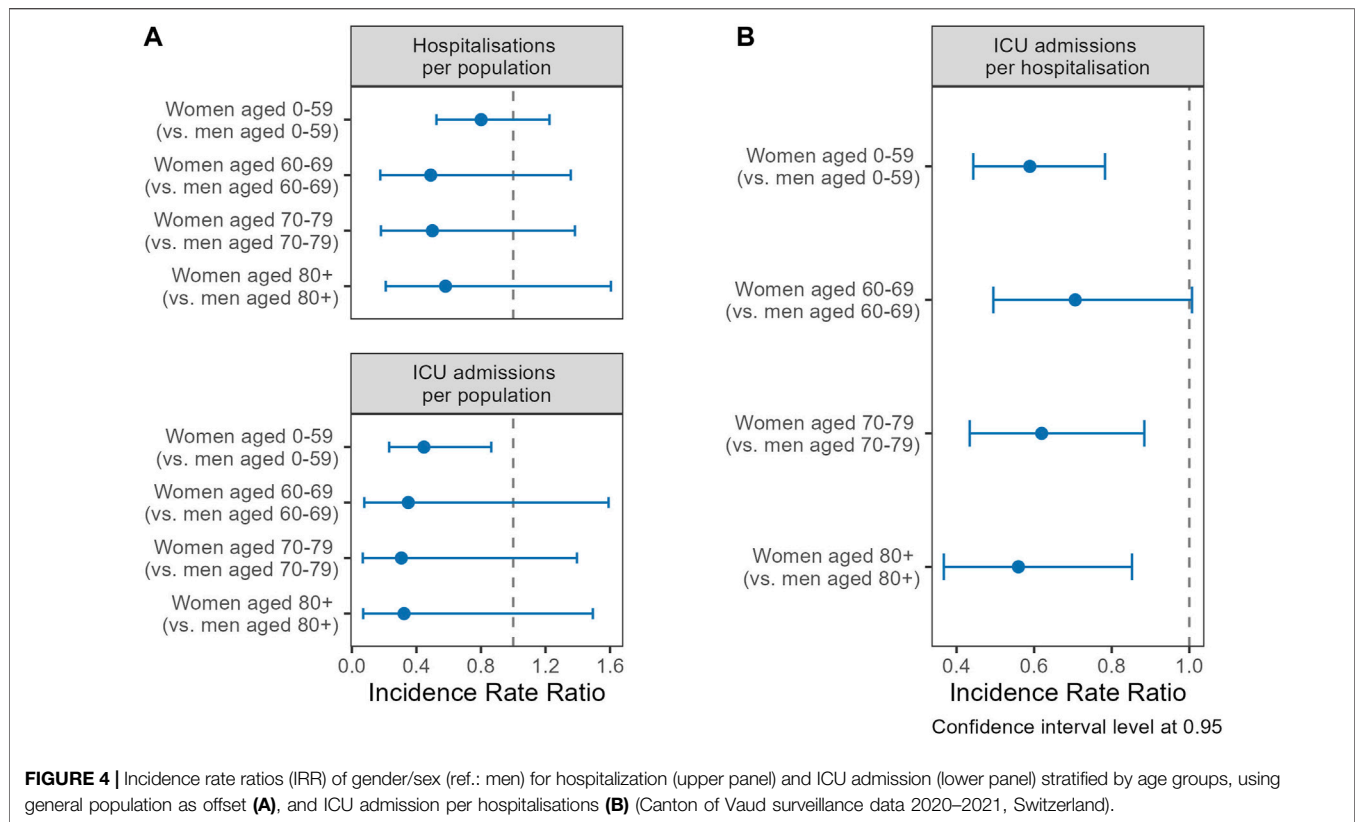
DISCUSSION

In the resident population of the Canton of Vaud, Switzerland, women contributed to a higher number of COVID-19 tests and positive tests than men, whereas more hospitalizations, ICU admissions, and deaths occurred among men. This finding underscores a pronounced gender/sex disparity in the

pandemic’s health impact, highlighting the need to explore underlying causes, such as potential biological differences, gender-specific behavioral patterns, and occupational exposures.

Individuals residing in the highest SEP neighborhoods underwent more COVID-19 testing than those in the lowest SEP areas, accompanied by a lower likelihood of testing positive and a reduced risk of mortality. These observations suggest significant socio-economic influences on health-related behaviors and resource accessibility and utilization. Notably, our intersectional analysis revealed that these disparities in testing and positivity rates are consistent across women and men within similar SEP quintiles. Moreover, the gender/sex disparities in mortality across SEP quintiles highlight the intricate interplay between socioeconomic factors and gender/sex, reinforcing the value of an intersectional approach in uncovering nuanced aspects of COVID-19 epidemiology.

Moreover, age-related variations in SARS-CoV-2 testing rates between women and men were evident in our data. Women aged



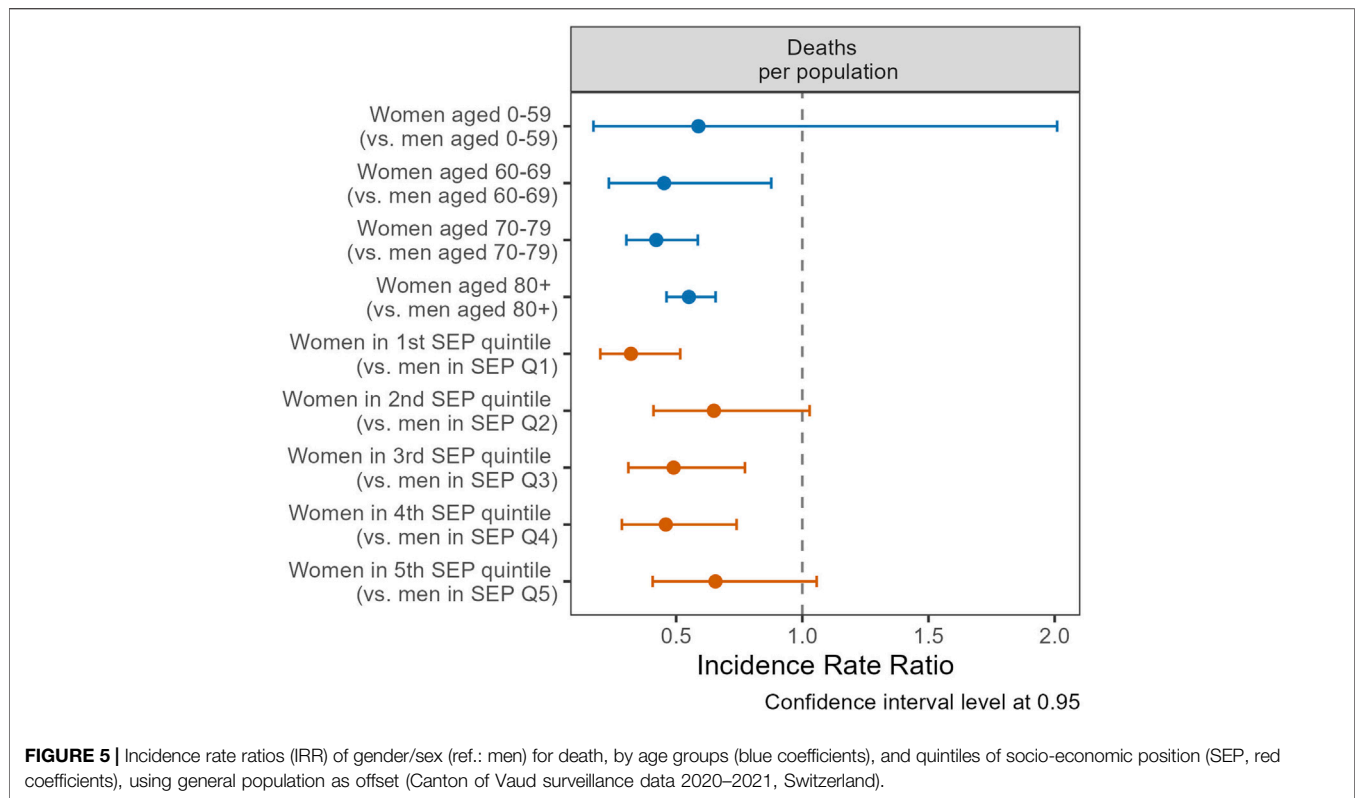
20–29 and 30–39 had higher testing rates than men in corresponding age groups, whereas this trend reversed in the younger (<20) and older age groups (60–69 and 70–79), echoing trends observed in other European countries [48]. These variations illustrate a complex relationship between age, gender/sex, and health-seeking behavior, calling for further investigations.

Age is a key factor in understanding COVID-19 gender/sex disparities. The influence of gender norms on health outcomes varies across the life course [19, 28, 31], and the disparities in testing rates between men and women across different age groups likely reflect the evolving societal roles and responsibilities [49]. In the 20 to 40 age range, where gender differences in testing were most pronounced, marked distinctions in family and employment domains are generally observed. Women are more likely to work in essential service sectors involving close contacts and limited telecommuting options, such as in service and healthcare jobs [49–52], which may account for their higher testing rates. Yet, this potential increased exposure did not translate into higher positivity rates when accounting for initial differences in testing, possibly attributable to greater adherence to health recommendations and protective measures among women compared to men [23, 26, 29, 51, 53]. Additionally, women in this age group often bear a disproportionate burden of unpaid care responsibilities, likely influencing their decisions regarding COVID-19 testing [29, 54]. The observed higher testing rates in men aged 60 and above may be attributed to the preferential ascertainment of severe cases [55]. Individuals who are

perceived as more likely to suffer from severe forms of infection—such as men, due to early reports of higher mortality rates—, are therefore tested more frequently. Although the ratios of positive tests were generally similar, women aged 60 and above were significantly less likely to test positive per test conducted compared to their male counterparts, highlighting possible differences in exposure.

Our study corroborates the well-established correlation between age and severe COVID-19 outcomes, with older age associated with an increased risk of hospitalization, ICU admissions, and mortality. Additionally, our data confirm that men face a higher risk of severe outcomes compared to women, aligning with previous research [56–58]. Notably, women under the age of 60 and those aged 70 and above had a reduced risk of ICU admissions when hospitalized, suggesting possible variations in immune system responses, prevalence of comorbidities, health-seeking behaviors, or differences in treatment approaches between women and men.

Our findings are consistent with existing literature on the link between SEP and COVID-19 outcomes [1–5], highlighting the increased vulnerability of individuals residing in low SEP neighborhoods. This vulnerability stems from a combination of factors, including limited access to healthcare, higher exposure risks due to living and working conditions, occupational hazards, and lifestyle habits, coupled with higher comorbidities rates [59]. Previous studies show that those in lower SEP areas experienced lower testing rates—particularly pronounced in the pandemic’s early stages—and faced elevated



rates of case incidence, hospitalizations, and mortality, a trend reported globally [1–5, 9, 60–62].

As extensively documented in the literature and corroborated by our findings, men experienced higher mortality rates related to COVID-19 compared to women [12, 13, 15–17, 23, 29, 63]. Our results outlined gender/sex disparities in mortality across SEP quintiles, particularly marked in the lowest SEP neighborhoods. This suggests that men from socioeconomically deprived backgrounds may encounter cumulative disadvantages that amplify their health vulnerabilities throughout their lives [64]. Biological sex-related factors are thought to play a significant role in men's increased vulnerability to COVID-19, possibly mediated through hormonal and immune response [29, 52]. Mortality is also influenced by gendered practices and societal norms such as expressions of masculinity, which intersect with various social determinants of health [52]. These include working in hazardous industries, engaging in risky health behaviors, and maintaining lifestyles that lead to higher prevalence of chronic diseases [26, 28, 29, 65, 66], which are more common in socioeconomically disadvantaged populations [60]. Beyond age, pre-existing medical conditions such as obesity, diabetes, or chronic respiratory and cardiovascular diseases have been described as major contributors to COVID-19 severity [67]. Comparatively riskier health behaviors among men, such as smoking, excessive alcohol consumption, and unhealthy diets, combined with societal norms that valorize toughness and discourage timely medical care, are thought to contribute to their health vulnerabilities [16, 19, 26]. Our findings contribute

to the growing body of intersectional studies on COVID-19 epidemiology, predominantly concentrated in the US and often emphasize ethnic/racial disparities alongside gender or sex. Notably, research within this domain showed that Black women experienced higher mortality rates than White men, while Black men had the highest mortality rates [68]. Another study exploring the combined effects of gender, SEP, and race/ethnicity revealed that, compared to high SEP White women, low SEP White men experienced a mortality rate 7.4 times higher, and 1.5 times higher compared to women in low SEP. Meanwhile, low SEP Hispanic men faced the most significant disparity, with mortality rates 27 times higher than those of high SEP White women [69].

Our study's strengths include the use of a neighborhood-based SEP indicator to capture potential individual and local-level effects on outcomes, and the minimization of selection bias by using comprehensive surveillance data for the entire Canton of Vaud population. Although Vaud was heavily impacted during the early stages of the pandemic, its diverse rural and urban population profiles provide valuable insights that may reflect broader trends in Switzerland, despite some regional variations. However, our analysis is limited by the absence of data on key individual-level factors such as migration background status and ethnicity. Incorporating these factors could greatly enrich our understanding, especially given that approximately one-third of Vaud's population in 2020 held non-Swiss nationality [41], and ethnic minorities faced higher exposure and vulnerability to COVID-19 [62, 68, 70]. Furthermore, disentangling the sources of disparities between gender and sex is methodologically

challenging when using administrative sex to investigate women's and men's health outcomes. Nevertheless, our intersectional approach facilitated the development of hypotheses about gendered mechanisms, which extend beyond the traditional biological interpretations common in biomedical research. Another limitation concerns hospitalization and ICU admission data, which may be subject to underreporting due to challenges associated with identifying primary causes of hospitalization, especially among older adults with comorbidities [71]. Moreover, deaths occurring outside clinical settings frequently remain untested, complicating their classification as COVID-19 related [71, 72]. Additionally, relying on residence location for the Swiss-SEP indicator may not accurately reflect an individual's lifelong SEP, and using a neighborhood-level indicator may also introduce inaccuracies when representing individual-level characteristics. These are common challenges associated with area-based indicators [73].

Conclusion

Our study within the Canton of Vaud highlighted the significant interplay of gender/sex, age, and SEP in shaping the epidemiology COVID-19. The intersectional analyses have revealed nuanced disparities, notably the increased risk of mortality in men, particularly those from lower SEP neighborhoods. While no substantial gender/sex differences in testing outcomes were observed across SEP quintiles, important age-related variation emerged, with young adult women experiencing higher testing rates. These findings underscore the importance of adopting intersectional approaches in both epidemiological research and public health strategy development. Such approaches are necessary for developing more effective and equitable health responses.

ETHICS STATEMENT

The studies involving humans were approved by Commission cantonale d'éthique de la recherche sur l'être humain du canton de Vaud (CER-VD). The studies were conducted in accordance with the local legislation and institutional requirements. The ethics committee/institutional review board waived the requirement of written informed consent for participation

REFERENCES

- Laajaj R, Webb D, Aristizabal D, Behrentz E, Bernal R, Buitrago G, et al. Understanding How Socioeconomic Inequalities Drive Inequalities in COVID-19 Infections. *Scientific Rep* (2022) 12(1):8269. doi:10.1038/s41598-022-11706-7
- Benita F, Rebollar-Ruelas L, Gaytán-Alfaro ED. What Have We Learned about Socioeconomic Inequalities in the Spread of COVID-19? A Systematic Review. *Sust Cities Soc* (2022) 86:104158. doi:10.1016/j.scs.2022.104158
- Vandentorren S, Smaïli S, Chatignoux E, Maurel M, Alleaume C, Neufcourt L, et al. The Effect of Social Deprivation on the Dynamic of SARS-CoV-2 Infection in France: A Population-Based Analysis. *The Lancet public health* (2022) 7(3):e240–e249. doi:10.1016/S2468-2667(22)00007-X
- Mongin D, Cullati S, Kelly-Irving M, Rosselet M, Regard S, Courvoisier DS, et al. Neighbourhood Socio-Economic Vulnerability and Access to COVID-19 Healthcare during the First Two Waves of the Pandemic in Geneva,

from the participants or the participants' legal guardians/next of kin because obtaining consent was considered infeasible or disproportionately burdensome. This exemption is in line with Article 34 of the Swiss Human Research Act (art. 34 LRH), which authorizes the exceptional reuse of health-related personal data for research purposes. For a subset of hospital admission data, a general consent process was in place. We excluded data from patients who objected to the reuse of their data.

AUTHOR CONTRIBUTIONS

YM, JS, and DA conceptualized and designed the research, supported by insights from all co-authors. Data acquisition and preparation were carried out by YM and VP. DA performed the statistical analysis, receiving critical and technical support from MA and JR. The initial draft of the manuscript was written by DA, which was then critically revised and edited by all other co-authors: YM, JS, CC, VP, MA, and JR. All authors contributed to the article and approved the submitted version.

FUNDING

This research received financial support from the “Pro-Femmes” scholarship program, granted by the Faculty of Biology and Medicine of the University of Lausanne.

CONFLICT OF INTEREST

The authors declare that they do not have any conflicts of interest.

SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: <https://www.sph-journal.org/articles/10.3389/ijph.2024.1607063/full#supplementary-material>

Switzerland: A Gender Perspective. *EClinicalMedicine* (2022) 46:101352. doi:10.1016/j.eclinm.2022.101352

- Riou J, Panczak R, Althaus CL, Junker C, Perisa D, Schneider K, et al. Socioeconomic Position and the COVID-19 Care Cascade From Testing to Mortality in Switzerland: A Population-Based Analysis. *The Lancet Public Health* (2021) 6(9):e683–e691. doi:10.1016/S2468-2667(21)00160-2
- Khazanchi R, Beiter ER, Gondi S, Beckman AL, Bilinski A, Ganguli I. County-Level Association of Social Vulnerability With COVID-19 Cases and Deaths in the USA. *J Gen Intern Med* (2020) 35(9):2784–7. doi:10.1007/s11606-020-05882-3
- Berchet C, Bijlholt J, Ando M. *Socio-Economic and Ethnic Health Inequalities in COVID-19 Outcomes Across OECD Countries* (2023).
- Jassat W, Ozougwu L, Munshi S, Mudara C, Vika C, Arendse T, et al. Corrigendum: The Intersection of Age, Sex, Race and Socio-Economic Status in COVID-19 Hospital Admissions and Deaths in South Africa. *South Afr J Sci* (2022) 118(5/6). doi:10.17159/sajs.2022/13323c
- Khan MM, Odoi A, Odoi EW. Geographic Disparities in COVID-19 Testing and Outcomes in Florida. *BMC Public Health* (2023) 23(1):79. doi:10.1186/s12889-022-14450-9

10. Tudor Hart J. The Inverse Care Law. *The Lancet* (1971) 297(7696):405–12. doi:10.1016/s0140-6736(71)92410-x
11. Corna LM. A Life Course Perspective on Socioeconomic Inequalities in Health: A Critical Review of Conceptual Frameworks. *Adv Life Course Res* (2013) 18(2):150–9. doi:10.1016/j.alcr.2013.01.002
12. Thompson K, Vassallo A, Finfer S, Woodward M. Renewed Rationale for Sex and Gender-Disaggregated Research: A COVID-19 Commentary Review. *Women's Health* (2022) 18:17455065221076738. doi:10.1177/17455065221076738
13. Gebhard C, Regitz-Zagrosek V, Neuhauser HK, Morgan R, Klein SL. Impact of Sex and Gender on COVID-19 Outcomes in Europe. *Biol Sex Differences* (2020) 11(1):29. doi:10.1186/s13293-020-00304-9
14. Scully EP, Schumock G, Fu M, Massaccesi G, Muschelli J, Betz J, et al. Sex and Gender Differences in Testing, Hospital Admission, Clinical Presentation, and Drivers of Severe Outcomes from COVID-19. *Open Forum Infect Dis* (2021) 8(9):ofab448. doi:10.1093/ofid/ofab448
15. Peckham H, de Grijter NM, Raine C, Radziszewska A, Ciurtin C, Wedderburn LR, et al. Male Sex Identified by Global COVID-19 Meta-Analysis as a Risk Factor for Death and ICU Admission. *Nat Commun* (2020) 11(1):6317–0. doi:10.1038/s41467-020-19741-6
16. Bamba C, Albani V, Franklin P. COVID-19 and the Gender Health Paradox. *Scand J Public Health* (2021) 49(1):17–26. doi:10.1177/1403494820975604
17. Sharma G, Volgman AS, Michos ED. Sex Differences in Mortality From COVID-19 Pandemic. *JACC: Case Rep* (2020) 2(9):1407–10. doi:10.1016/j.jaccas.2020.04.027
18. Bauer GR. Sex and Gender Multidimensionality in Epidemiologic Research. *Am J Epidemiol* (2023) 192(1):122–32. doi:10.1093/aje/kwac173
19. Heise L, Greene ME, Opper N, Stavropoulou M, Harper C, Nascimento M, et al. Gender Inequality and Restrictive Gender Norms: Framing the Challenges to Health. *The Lancet* (2019) 393(10189):2440–54. doi:10.1016/S0140-6736(19)30652-X
20. Kaufman MR, Eschliman EL, Karver TS. Differentiating Sex and Gender in Health Research to Achieve Gender Equity. *Bull World Health Organ* (2023) 101(10):666–71. doi:10.2471/BLT.22.289310
21. Shannon G, Jansen M, Williams K, Cáceres C, Motta A, Odhiambo A, et al. Gender Equality In Science, Medicine, and Global Health: Where Are We at and Why Does It Matter? *Lancet* (2019) 393(10171):560–9. doi:10.1016/S0140-6736(18)33135-0
22. Laster Pirtle WN, Wright T. Structural Gendered Racism Revealed in Pandemic Times: Intersectional Approaches to Understanding Race and Gender Health Inequities in COVID-19. *Gend Soc* (2021) 35(2):168–79. doi:10.1177/08912432211001302
23. Danielsen AC, Lee KM, Boulicault M, Rushovich T, Gompers A, Tarrant A, et al. Sex Disparities in COVID-19 Outcomes in the United States: Quantifying and Contextualizing Variation. *Soc Sci Med* (2022) 294:114716. doi:10.1016/j.socscimed.2022.114716
24. Clark C, Davila A, Regis M, Kraus S. Predictors of COVID-19 Voluntary Compliance Behaviors: An International Investigation. *Glob Transitions* (2020) 2:76–82. doi:10.1016/j.glt.2020.06.003
25. Wenham C, Smith J, Morgan R, Gender and COVID-19 Working Group. COVID-19: The Gendered Impacts of the Outbreak. *The Lancet* (2020) 395(10227):846–8. doi:10.1016/S0140-6736(20)30526-2
26. Ya'qoub L, Elgendy IY, Pepine CJ. Sex and Gender Differences in COVID-19: More to Be Learned!. *Am Heart J Plus: Cardiol Res Pract* (2021) 3:100011. doi:10.1016/j.ahjo.2021.100011
27. Groban L, Wang H, Sun X, Ahmad S, Ferrario CM. Is Sex a Determinant of COVID-19 Infection? Truth or Myth? *Curr Hypertens Rep* (2020) 22(9):62. doi:10.1007/s11906-020-01073-x
28. Taslem Mouroso J, Anwar S, Hosen MJ. The Sex and Gender Dimensions of COVID-19: A Narrative Review of the Potential Underlying Factors. *Infect Genet Evol* (2022) 103:105338. doi:10.1016/j.meegid.2022.105338
29. Diaz-Rodriguez N, Binkyte R, Bakkali W, Bookseller S, Tubaro P, Bacevicius A, et al. Gender and Sex Bias in COVID-19 Epidemiological Data Through the Lens of Causality. *Inf Process Manag* (2023) 60(3):103276. doi:10.1016/j.ipm.2023.103276
30. Harris MA, Wagner DV, Heywood M, Hoehn D, Bahia H, Spiro K. Youth Repeatedly Hospitalized for DKA: Proof of Concept for Novel Interventions in Children's Healthcare (NICH). *Diabetes Care* (2014) 37:doi:10.2337/dc13-2232
31. Notarte KI, de Oliveira MHS, Peligro PJ, Velasco JV, Macaranas I, Ver AT, et al. Age, Sex and Previous Comorbidities as Risk Factors Not Associated with SARS-CoV-2 Infection for Long COVID-19: A Systematic Review and Meta-Analysis. *J Clin Med* (2022) 11(24):7314. doi:10.3390/jcm11247314
32. Hammarström A, Johansson K, Annandale E, Ahlgren C, Aléx L, Christianson M, et al. Central Gender Theoretical Concepts in Health Research: The State of the Art. *J Epidemiol Community Health* (2014) 68(2):185–90. doi:10.1136/jech-2013-202572
33. Bauer GR, Churchill SM, Mahendran M, Walwyn C, Lizotte D, Villa-Rueda AA. Intersectionality in Quantitative Research: A Systematic Review of its Emergence and Applications of Theory and Methods. *SSM Popul Health* (2021) 14:100798. doi:10.1016/j.ssmph.2021.100798
34. Olanlesi-Aliu A, Tulli M, Kemei J, Bonifacio G, Reif LC, Cardo V, et al. A Scoping Review on the Operationalization of Intersectional Health Research Methods in Studies Related to the COVID-19 Pandemic. *Int J Qual Stud Health Well-being* (2024) 19(1):2302305. doi:10.1080/17482631.2024.2302305
35. Collins PH. Intersectionality's Definitional Dilemmas. *Annu Rev Sociol* (2015) 41(1):1–20. doi:10.1146/annurev-soc-073014-112142
36. Molenaar J. *Using an Intersectional Lens to Understand the Unequal Impact of the COVID-19 Pandemic* (2021). Bi-monthly report 5 September 2021.
37. Bauer GR. Incorporating Intersectionality Theory into Population Health Research Methodology: Challenges and the Potential to Advance Health Equity. *Soc Sci Med* (2014) 110:10–7. doi:10.1016/j.socscimed.2014.03.022
38. Giachino M, Valera CBG, Rodriguez Velásquez S, Dohrendorf-Wyss MA, Rozanova L, Flahault A. Understanding the Dynamics of the COVID-19 Pandemic: A Real-Time Analysis of Switzerland's First Wave. *Int J Environ Res Public Health* (2020) 17(23):8825. doi:10.3390/ijerph17238825
39. Bag.Admin. Office fédéral de la santé publique, Commission fédérale pour les vaccinations. COVID-19: Stratégie de vaccination (2022). Available from: <https://www.bag.admin.ch/dam/bag/fr/dokumente/mt/k-und-i/aktuelle-ausbrueche-pandemien/2019-nCoV/impfstrategie-bag-ekif.pdf>. download.pdf/strategie-de-vaccination-covid-19-ofsp-ekif.pdf (Accessed November 29, 2022).
40. EPICoVID. Direction Générale de la santé (DGS), Office du médecin cantonal. COVID-19: Point épidémiologique – Situation au (2021). Available from: https://infosan.vd.ch/fileadmin/2-PUBLICATIONS/SANTE_POPULATION/SSP_20210628_COVID_Bulletin_hebdomadaire_epidemiopdf (Accessed June 28, 2021).
41. Statistique Vaud. Population résidente permanente par âge exact_ sexe et origine_ Vaud_ 2017-2020 (2020). Available from: <https://www.vd.ch/themes/etat-droit-finances/statistique/statistiques-par-domaine/01-population/etat-et-structure-de-la-population> (Accessed April 20, 2022).
42. Panczak R, Berlin C, Voorpostel M, Zwahlen M, Egger M. The Swiss Neighbourhood Index of Socioeconomic Position: Update and Re-Validation. *Swiss Med Weekly* (2023) 153(1):40028. doi:10.57187/smw.2023.40028
43. Panczak R, Galobardes B, Voorpostel M, Spoerri A, Zwahlen M, Egger M, et al. A Swiss Neighbourhood Index of Socioeconomic Position: Development and Association with Mortality. *J Epidemiol Community Health* (2012) 66(12):1129–36. doi:10.1136/jech-2011-200699
44. R Core Team. R: A Language and Environment for Statistical Computing. In: *Computing RfS*. Vienna, Austria: Wiley (2022).
45. Ripley B, Venables B. *Modern Applied Statistics with S*. 4th ed. New York: Springer (2022).
46. Heidari S, Babor TF, De Castro P, Tort S, Curno M. Sex and Gender Equity in Research: Rationale for the SAGER Guidelines and Recommended Use. *Res Integr Peer Rev* (2016) 1(1):2. doi:10.1186/s41073-016-0007-6
47. Fausto-Sterling A. Gender/Sex, Sexual Orientation, and Identity Are in the Body: How Did They Get There? *J Sex Res* (2019) 56(4-5):529–55. doi:10.1080/00224499.2019.1581883
48. Sobotka T, Brzozowska Z, Mutarak R, Zeman K, di Lego V. Age, Gender and COVID-19 Infections. medRxiv. (2020):
49. Yavorsky JE, Qian Y, Sargent AC. The Gendered Pandemic: The Implications of COVID-19 for Work and Family. *Sociol Compass* (2021) 15(6):e12881. doi:10.1111/soc4.12881

50. Kley S, Reimer T. Exploring the Gender Gap in Teleworking from Home. The Roles of Worker's Characteristics, Occupational Positions and Gender Equality in Europe. *Soc Indicators Res* (2023) 168(1):185–206. doi:10.1007/s11205-023-03133-6
51. Sant FC, Fischer FB, Monzón Llamas L, Tavakkoli M, Cobos Muñoz D, Antillon M. Did COVID-19 Policies Have the Same Effect on COVID-19 Incidence Among Women and Men? Evidence from Spain and Switzerland. *Int J Public Health* (2022) 67:1604994. doi:10.3389/ijph.2022.1604994
52. Morgan R, Baker P, Griffith DM, Klein SL, Logie CH, Mwiine AA, et al. Beyond a Zero-Sum Game: How Does the Impact of COVID-19 Vary by Gender? *Front Sociol* (2021) 6:126. doi:10.3389/fsoc.2021.650729
53. Galasso V, Pons V, Profeta P, Becher M, Brouard S, Foucault M. Gender Differences in COVID-19 Attitudes and Behavior: Panel Evidence from Eight Countries. *Proc Natl Acad Sci* (2020) 117(44):27285–91. doi:10.1073/pnas.2012520117
54. Bühler N, Pralong M, Rawlinson C, Gonseth S, D'Acremont V, Bochud M, et al. Caring During COVID-19: Reconfigurations of Gender and Family Relations During the Pandemic in Switzerland. *Front Sociol* (2021) 6:737619. doi:10.3389/fsoc.2021.737619
55. Lipsitch M, Donnelly CA, Fraser C, Blake IM, Cori A, Dorigatti I, et al. Potential Biases in Estimating Absolute and Relative Case-Fatality Risks During Outbreaks. *Plos Negl Trop Dis* (2015) 9(7):e0003846. doi:10.1371/journal.pntd.0003846
56. Hawkes S, Pantazis A, Purdie A, Gautam A, Kitiwuwa-Muyingo S, Buse K, et al. Sex-Disaggregated Data Matters: Tracking the Impact of COVID-19 on the Health of Women and Men. *Econ Polit (Bologna)* (2022) 39(1):55–73. doi:10.1007/s40888-021-00254-4
57. Pivonello R, Auriemma RS, Pivonello C, Isidori AM, Corona G, Colao A, et al. Sex Disparities in COVID-19 Severity and Outcome: Are Men Weaker or Women Stronger? *Neuroendocrinology* (2021) 111(11):1066–85. doi:10.1159/000513346
58. Scully EP, Schumock G, Fu M, Massaccesi G, Muschelli J, Betz J. Sex and Gender Differences in Testing, Hospital Admission, Clinical Presentation, and Drivers of Severe Outcomes from COVID-19. *Open Forum Infectious Diseases*. Oxford: Oxford University Press US (2021).
59. Acebillo-Baqué M, Maestriperi L. Intersectionality Theory and Its Application in the COVID-19 Pandemics. In: Liamputtong P, editor. *Handbook of Social Sciences and Global Public Health*. Cham: Springer International Publishing (2023). p. 1189–211.
60. McGowan VJ, Bamba C. COVID-19 Mortality and Deprivation: Pandemic, Syndemic, and Endemic Health Inequalities. *The Lancet Public Health* (2022) 7(11):e966–e975. doi:10.1016/S2468-2667(22)00223-7
61. Bamba C. Pandemic Inequalities: Emerging Infectious Diseases and Health Equity. *Int J Equity Health* (2022) 21(1):6. doi:10.1186/s12939-021-01611-2
62. Khanijahani A, Iezadi S, Gholipour K, Azami-Aghdash S, Naghibi D. A Systematic Review of Racial/Ethnic and Socioeconomic Disparities in COVID-19. *Int J Equity Health* (2021) 20(1):248. doi:10.1186/s12939-021-01582-4
63. Kim H, Fox AM, Kim Y, Kim R, Bae G, Kang M. Is the Male Disadvantage Real? Cross-National Variations in Sex Gaps in COVID-19 Incidence and Mortality. *Glob Public Health* (2021) 16(12):1793–803. doi:10.1080/17441692.2021.1981972
64. Turrell G, Lynch JW, Leite C, Raghunathan T, Kaplan GA. Socioeconomic Disadvantage in Childhood and Across the Life Course and All-Cause Mortality and Physical Function in Adulthood: Evidence From the Alameda County Study. *J Epidemiol Community Health* (2007) 61(8):723–30. doi:10.1136/jech.2006.050609
65. Anderegg N, Panczak R, Egger M, Low N, Riou J. Survival Among People Hospitalized With COVID-19 in Switzerland: A Nationwide Population-Based Analysis. *BMC Med* (2022) 20(1):164. doi:10.1186/s12916-022-02364-7
66. Nikoloski Z, Alqunaibet AM, Alfawaz RA, Almudarra SS, Herbst CH, El-Saharty S, et al. Covid-19 and Non-Communicable Diseases: Evidence From a Systematic Literature Review. *BMC Public Health* (2021) 21(1):1068. doi:10.1186/s12889-021-11116-w
67. Agodi A, Maugeri A, Favara G, Magnano San Lio R, Puglisi M, Sinatra D, et al. Gender Differences in Comorbidities of Patients With COVID-19: An Italian Local Register-Based Analysis. *Heliyon* (2023) 9(7):e18109. doi:10.1016/j.heliyon.2023.e18109
68. Rushovich T, Boulicault M, Chen JT, Danielsen AC, Tarrant A, Richardson SS, et al. Sex Disparities in COVID-19 Mortality Vary Across US Racial Groups. *J Gen Intern Med* (2021) 36(6):1696–701. doi:10.1007/s11606-021-06699-4
69. Pathak EB, Menard JM, Garcia RB, Salemi JL. Joint Effects of Socioeconomic Position, Race/Ethnicity, and Gender on COVID-19 Mortality Among Working-Age Adults in the United States. *Int J Environ Res Public Health* (2022) 19(9):5479. doi:10.3390/ijerph19095479
70. Irizar P, Pan D, Kapadia D, Bécares L, Sze S, Taylor H, et al. Ethnic Inequalities in COVID-19 Infection, Hospitalisation, Intensive Care Admission, and Death: A Global Systematic Review and Meta-Analysis of Over 200 Million Study Participants. *eClinicalMedicine* (2023) 57.
71. Akter S. The Gender Gap in COVID-19 Mortality in the United States. *Feminist Econ* (2021) 27(1-2):30–47. doi:10.1080/13545701.2020.1829673
72. Riou J, Hauser A, Fesser A, Althaus CL, Egger M, Konstantinoudis G. Direct and Indirect Effects of the COVID-19 Pandemic on Mortality in Switzerland. *Nat Commun* (2023) 14(1):90. doi:10.1038/s41467-022-35770-9
73. Bryere J, Pornet C, Copin N, Launay L, Gusto G, Grosclaude P, et al. Assessment of the Ecological Bias of Seven Aggregate Social Deprivation Indices. *BMC Public Health* (2017) 17(1):86. doi:10.1186/s12889-016-4007-8

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Gender/Sex Disparities in the COVID-19 Cascade from Testing to Mortality: An Intersectional Analysis of Swiss Surveillance Data

Supplementary Table 1: Case definitions and notification process of outcomes, cantonal surveillance data of canton of Vaud, 2020-2021, Switzerland

Outcome	Case definition
Total tests	Tests reported to the Federal Office of Public Health (FOPH) as per FOPH's declaration criteria, as defined in the Ordinance of the Federal Department of Inner Affairs on declarations of observations related to transmissible diseases in humans (Ordonnance sur la déclaration d'observations en rapport avec les maladies transmissibles de l'homme 818.101.126) and attributed to the canton of Vaud.
Positive tests	Laboratory-confirmed tests (positive RT-PCR test or rapid antigen test) among total tests notified to the FOPH.
Hospitalisations and Intensive Care Unit (ICU) admissions	<p>Direct reporting from hospitals and clinics to the Office du Médecin Cantonal (OMC) through the EPICOVID cantonal surveillance system. Included records cover hospital stays in a healthcare facility of the canton of Vaud or in an intercantonal facility. Up to April 15th, 2020, any stay of a patient with a positive SARS-CoV-2 RT-PCR test was included. From April 15th, 2020, inclusion criteria were refined to any stay of a patient with at least one positive RT-PCR test for SARS-CoV-2 and a viral load > 1000 copies during the hospital stay. Admission dates correspond to the entire stay, independent of the date of the viral load test or end of isolation. Each new stay is re-analyzed and included according to these definitions.</p> <p>Excluded records:</p> <ul style="list-style-type: none"> - Hospital stays below 24hrs - Stays in Centre de Traitement et Réadaptation (CTR) or in mental health facility - For patients admitted to the Centre Hospitalier Universitaire Vaudois (CHUV – Lausanne University Hospital), those who opted out of the use of their data for research purposes via the “consentement general” were excluded (N=510) <p>Note: Inclusion criteria vary based on the possibility of automated data extraction from hospitals. Criteria presented above correspond to those of CHUV. For other hospitals, inclusion criteria align with mandatory reporting criteria formulated by the Swiss Federal Office of Public Health (OFSP) as follows (as of 18.05.2020):</p> <p>Patient hospitalized for more than 24 hours who presents:</p> <ul style="list-style-type: none"> • Clinical criteria compatible with COVID-19 and a positive SARS-CoV-2 PCR test, or • Clinical criteria and CT-Scan imaging compatible with COVID-19 and a negative SARS-CoV-2 PCR test without any other known aetiology, or • Clinical and epidemiological criteria and a negative SARS-CoV-2 test without any other known aetiology.
Deaths	Deaths among probable or confirmed COVID-19 cases reported to the cantonal authorities.

Supplementary Section 2: Geocoding procedures and socio-economic position (SEP) attribution

Residential addresses were geocoded using publicly available data of the Federal Register of Buildings and Dwellings (under ORegBL ; RS 431.841) obtained from the Federal Office of Statistics (1). This register contains all residential addresses of Switzerland, at the cantonal level, with corresponding geographic coordinates based on the Swiss LV95 reference system. The process of geocoding and SEP attribution is detailed in figure 2B, next page. Geocoding was performed using both perfect and fuzzy

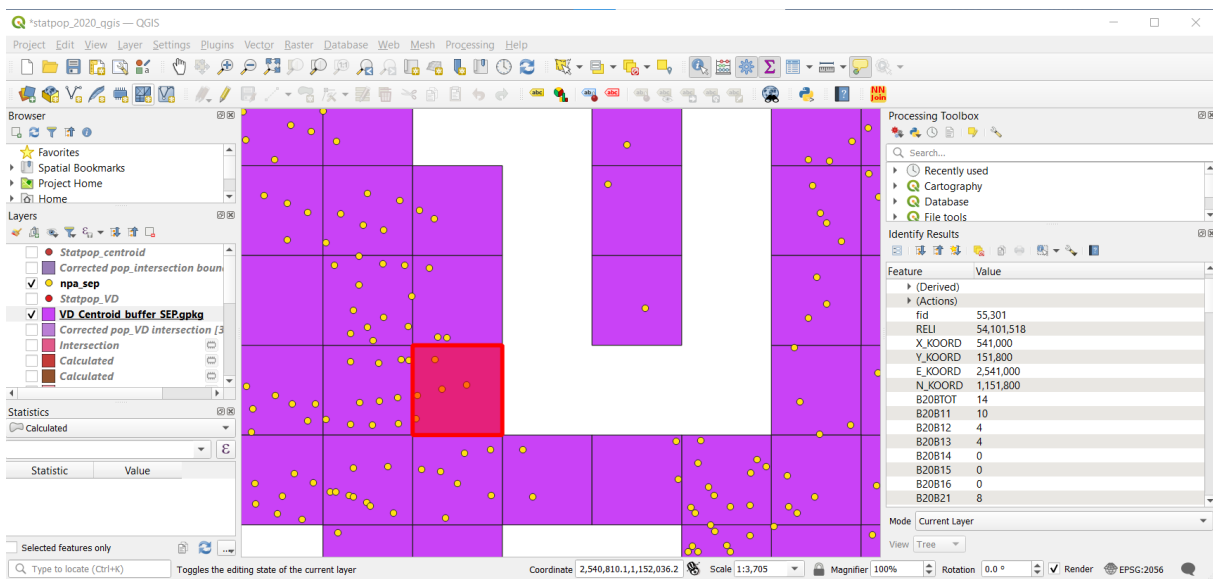
Supplementary Material | Gender/Sex Disparities in the COVID-19 Cascade from Testing to Mortality: An Intersectional Analysis of Swiss Surveillance Data

matching methods. Fuzzy matching involves comparing and matching text strings based on their similarity, taking into account variations in word order, spelling, and other factors. The *fuzzyjoin* package in R (2) was used, and distance between sequences of addresses calculated with the *Jaro-Winkler similarity method*, which provides a score between 0 (perfect similarity) and 1 (total dissimilarity). To derive the socio-economic position (SEP) score, ranging from 0 to 100, the residential coordinates in the notifications were matched with the nearest SEP neighborhood using the *SF* package of R (3).

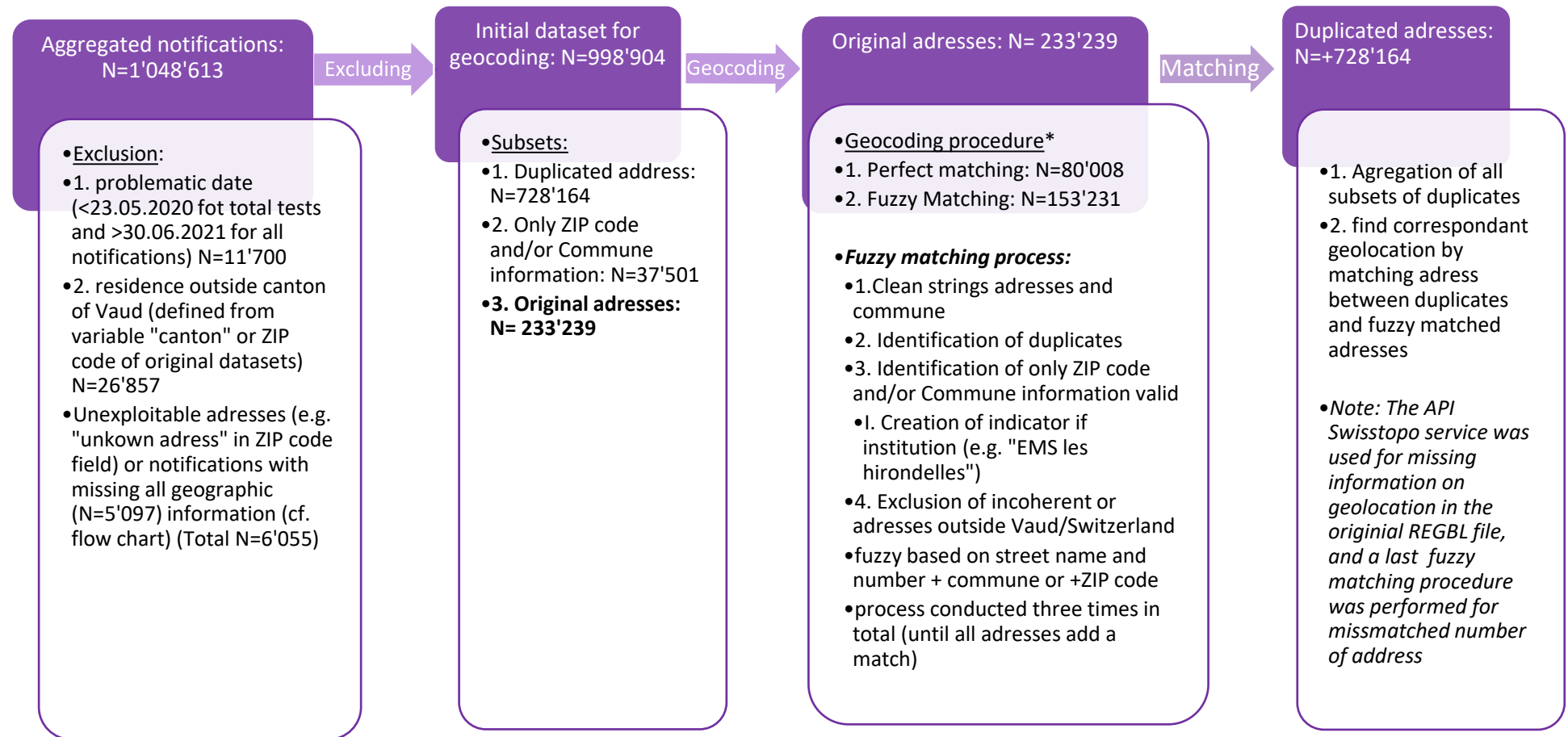
The spatial distribution of the Swiss-SEP index within the Vaud population was estimated to define the SEP cut-off quintiles and to establish the population dataset employed as the denominator for our analyses. This estimation employed Statpop hectometric population data aggregated by age groups (5 years age-range) and sex/gender categories (women, men) (4). All geographic information system processes were performed using the software QGIS (5).

To determine the distribution of SEP within the canton of Vaud's population, hectare-level aggregated data was utilized. For each hectare (depicted as purple squares in the illustration below), SEP values were retrieved or calculated – using the mean of SEP values when multiple SEP neighborhoods fell within a hectare – depicted as yellow dots in Figure 2A. These SEP values were then attributed to the respective population residing within each hectare. SEP values of the population were categorized into five groups using quintiles as cut-offs, ranging from one (lowest) to five (highest). These quintile cut-offs were then used to classify SEP scores of data notifications and assign them to the corresponding quintiles.

Supplementary Figure 2A: Illustration of hectometric population data (purple squares) and SEP neighborhoods (yellow dots, among the 115'596 SEP neighborhoods in Vaud) on QGIS



Supplementary Figure 2B: Flow chart of geocoding procedure of the residential addresses of notifications



Finally, to establish the necessary dataset for denominator population, we analyzed the distribution of SEP quintiles across age and sex/gender categories within the hectare-level aggregated data and then applied this distribution to the Vaud population dataset, which was disaggregated by age groups and sex categories. This data was obtained from the *Statistique Vaud* (6), specifically the “Permanent resident population by exact age, gender, and origin, Vaud” dataset, as of December 31st, 2020. This dataset, derived from cantonal registers of population, is notably more precise than the GEOSTAT Statpop data. However, in this latter dataset, the cantonal boundaries do not align with the hectometric grid, and the adjustment of population size by proportional coverage within the cantonal boundaries is less than ideal. Additionally, for confidentiality reasons, the GEOSTAT dataset employs a uniform censoring approach for population subgroups with fewer than three individuals, assigning them a standardized value of three. While this practice aims to protect anonymity, it also leads to overestimations in population size.

Supplementary Section 3: the Swiss socio-economic position (Swiss-SEP)

The Swiss-SEP is an area-based socio-economic position index centered on residential buildings, encompassing overlapping boundaries. The Swiss-SEP considers neighborhood information based on an average of 50 households (7) and is widely used in social and health research to investigate the associations between social and economic position and health outcomes, as seen in recent studies (8-10). The Swiss-SEP is constructed using principal component analysis, deriving information from the 2000 population census on four indicators as proxies for SEP: median rent per square meter (income proxy), proportion of households led by individuals with a primary education or less (education proxy), proportion of households with individuals in manual or unskilled jobs (occupation proxy), and the average number of persons per room (crowding proxy). The index scores range from 0 to 100, with higher values indicating higher SEP. The Swiss-SEP scores range from 1 to 100, with higher scores indicating a higher SEP. In our study, we utilized the latest updated hybrid version, Swiss-SEP 3 (11), which incorporates information from the 2012 to 2015 annual micro-censuses and updated values for neighborhood centered on buildings constructed after 2000.

Within the geographical boundaries of Vaud, we identified a total of 115'596 SEP neighborhoods. We then matched the residential coordinates of each notification with the nearest SEP neighborhood. SEP index scores were subsequently aggregated into five categories using quintiles as cutoff points, ranging from one (lowest) to five (highest). In cases involving non-residential addresses, such as those corresponding to schools or nursing homes, as well as incomplete addresses, we used the mean SEP of the respective ZIP code area as a proxy for SEP. However, it's noteworthy that notifications related to hospitalization stays and ICU admissions lacked SEP attribution, as these notifications only contained ZIP code information. The reliance on mean SEP from ZIP code areas significantly influenced the distribution within quintiles, leading to a skewed categorization in the lowest and highest quintiles. Therefore, SEP information for hospitalizations and ICU admission outcomes was excluded from our analyses. Given that a significant proportion of death notifications could not be accurately geocoded, analyses using SEP quintiles were likewise restricted to geocoded notifications.

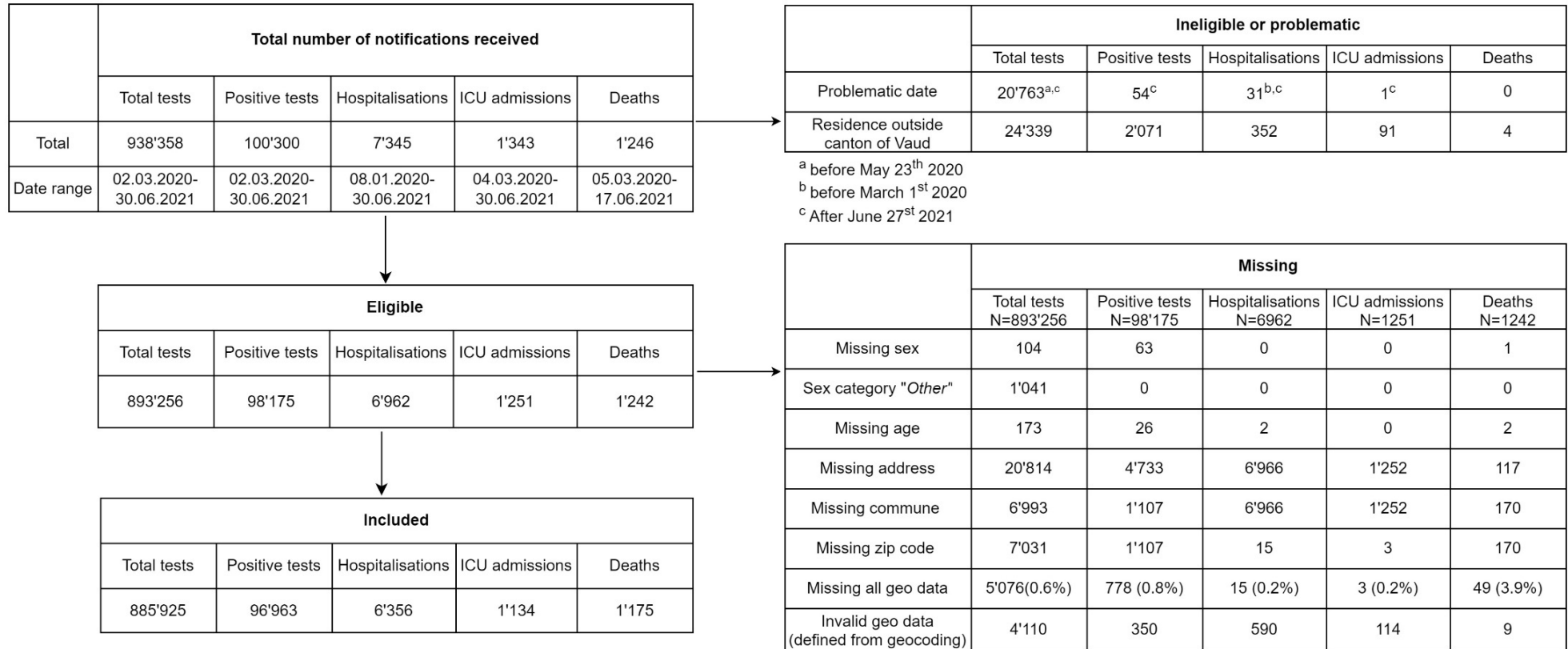
Supplementary Table 4: Geocoding status of notifications, by COVID-19 outcome

	Tests N (%) N= 885'925	Positive tests N (%) N=96'963	Hospitalisations N (%) N=6'356	ICU admissions N (%) N=1'134	Deaths N (%) N=1'175
Geocoded	834'912 (94.2%)	89'183 (91.9%)	0 (0%)	0 (0%)	713 (60.7%)
Incomplete address	51'013 (4.6%)	7'780 (6.9%)	6'359 (100%)	1'134 (100%)	462 (21.6%)
Non-residential location	10'779 (1.2%)	1'157 (1.2%)	NA	NA	207 (17.7%)

For both total tests and positive tests, approximately 4.6% and 8%, respectively, could not be geocoded and were therefore attributed to the mean socioeconomic position (SEP) of the respective ZIP code area. Moreover, approximately 1.2% of these notifications corresponded to non-residential addresses (i.e., institutions), receiving a similar SEP attribution based on the ZIP code area.

Regarding deaths, 60.7% of notifications included a complete residential address and were successfully geocoded. Women had a higher proportion of geocodes corresponding to institutional locations (22%), predominantly nursing homes, compared to men (14%).

Supplementary Figure 5: Flow chart of included notifications



Supplementary Table 6: Incidence rate ratios of total tests, positive tests, hospitalisations, ICU admissions and deaths, negative binomial regression models without interaction terms

	<i>Outcomes</i>					
	Total tests per population	Positive tests per population	Positive tests per test	Hospitalisations per population	ICU admissions per population	Deaths per population
	IRR (95% CI)	IRR (95% CI)	IRR (95% CI)	IRR (95% CI)	IRR (95% CI)	IRR (95% CI)
(ref. Men)						
Women	1.010 (0.98-1.041)	0.989 (0.953-1.026)	0.970* (0.945-0.996)	0.692* (0.488-0.981)	0.399*** (0.235-0.678)	0.510*** (0.427-0.608)
(ref. Age 50-59)						
Age 0-9	0.250***(0.235-0.267)	0.110***(0.101-0.121)	0.439***(0.408-0.473)			
Age 10-19	0.977 (0.917-1.041)	0.776***(0.719-0.837)	0.796***(0.755-0.838)			
Age 20-29	1.438***(1.349-1.532)	1.185***(1.1-1.277)	0.823***(0.783-0.866)			
Age 30-39	1.410***(1.323-1.502)	1.059 (0.983-1.141)	0.752***(0.714-0.791)			
Age 40-49	1.131***(1.061-1.205)	1.010 (0.937-1.088)	0.893***(0.848-0.939)	(ref. Age 0-59)	(ref. Age 0-59)	(ref. Age 0-59)
Age 60-69	0.881***(0.827-0.939)	0.861***(0.798-0.929)	0.976 (0.925-1.028)	4.833*** (2.855-8.786)	8.589*** (3.999-21.478)	29.818*** (15.785-61.282)
Age 70-79	0.755***(0.709-0.805)	0.764***(0.707-0.825)	1.007 (0.954-1.064)	8.103*** (4.791-14.72)	11.594*** (5.4-28.986)	154.327*** (87.466-303.293)
Age 80+	1.111** (1.042-1.185)	1.318***(1.22-1.424)	1.185***(1.122-1.251)	18.377*** (10.883-33.345)	9.612*** (4.439-24.171)	843.06*** (485.617-1638.875)
(ref. Q1) - lowest						
2nd SEP quintile	1.093*** (1.042-1.147)	1.033 (0.975-1.095)	0.942** (0.904-0.981)	NA	NA	0.954 (0.725-1.254)
3rd SEP quintile	1.166*** (1.111-1.223)	1.057 (0.998-1.121)	0.904*** (0.868-0.942)	NA	NA	0.982 (0.749-1.289)
4th SEP quintile	1.241*** (1.183-1.301)	0.970 (0.915-1.029)	0.775*** (0.744-0.808)	NA	NA	0.816 (0.617-1.08)
5th SEP quintile - highest	1.247*** (1.189-1.308)	0.897*** (0.846-0.951)	0.712*** (0.683-0.742)	NA	NA	0.714* (0.539-0.946)

Note: SEP= Socioeconomic position; NA= Not Applicable; p-value levels: * p<0.05; ** p<0.01; *** p<0.001
 Regression models: glm.nb(n_outcome~SEX+AGE+SEP+offset(log(population))

Supplementary Table 7: Sensitivity analysis - incidence rate ratios of death notifications using the general population as denominator, employing negative binomial regression models without interaction terms

As a significant proportion of death notifications could not be accurately geocoded due to incomplete addresses, which either contained only ZIP code information or corresponded to institutional locations, we conducted a sensitivity analysis by including notifications with imprecise SEP attribution and relying on the mean SEP of the ZIP code area.

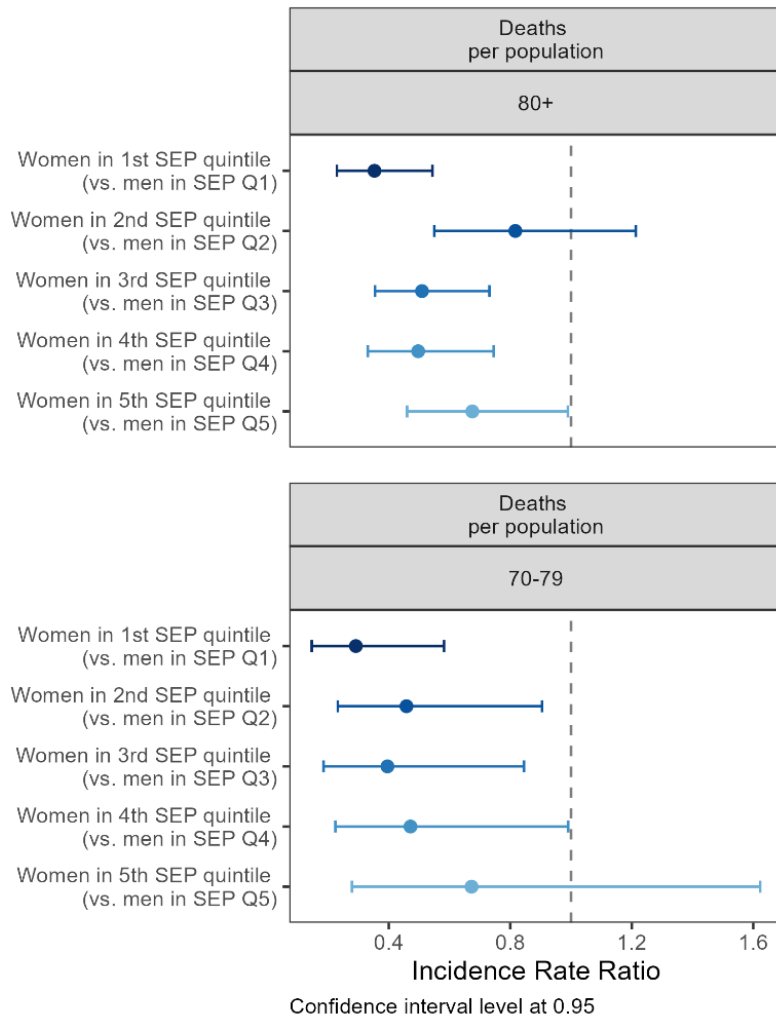
	(1)	(2)	(3)
	Geocoded death notifications, per population (N=713)	Death notifications outside institutions, per population (N=968)	All deaths notifications, per population (N=1175)
<i>Variables</i>	IRR (95% CI)	IRR (95% CI)	IRR (95% CI)
(ref. Men)			
Women	0.51*** (0.427-0.608)	0.514*** (0.439-0.601)	0.569*** (0.499-0.648)
(ref. Age 50-59)			
Age 60-69	29.818*** (15.785-61.282)	26.49*** (15.65-47.215)	25.813*** (15.664-44.534)
Age 70-79	154.327*** (87.466-303.293)	124.924*** (78.104-213.853)	129.063*** (82.797-213.874)
Age 80+	843.06*** (485.617-1638.875)	754.753*** (479.575-1276.94)	824.332*** (536.944-1350.711)
(ref. Q1)			
2nd SEP quintile	0.954 (0.725-1.254)	1.298* (1.02-1.656)	1.426** (1.153-1.766)
3rd SEP quintile	0.982 (0.749-1.289)	1.218 (0.955-1.557)	1.643*** (1.335-2.027)
4th SEP quintile	0.816 (0.617-1.08)	0.93 (0.721-1.2)	1.061 (0.849-1.327)
5th SEP quintile - highest	0.714* (0.539-0.946)	0.711* (0.547-0.925)	0.762* (0.603-0.962)

Note: Death notifications excluding those with incomplete residential addresses (column 1); Excluding institution-related addresses (column 2); Total death notifications (column 3); p-value levels: *p<0.05; **p<0.01; ***p<0.001

Excluding notifications with less precise SEP attribution had an impact on the regression results. When all notifications were included (column 3), our analysis suggested that individuals in the second and third quintiles had higher mortality rates compared to individual living in the lowest SEP areas. We currently lack a definitive explanation for these findings. Conversely, when restricting the analysis to notifications that could be accurately geocoded (column 1), with more precise SEP attribution, these associations became statistically non-significant. Hence, we opted to focus our analysis on death notifications with geocoded residential addresses for accurate SEP attribution.

However, the association between sex/gender and death in the highest quintile remained consistent across all analyses. Individuals in the highest SEP quintile were significantly less likely to die compared to those in the lowest quintile.

Supplementary Figure 8: Incidence rate ratios of death notifications, stratified for age groups 70-79 and 80+, using Poisson regression models



Notes: This figure illustrates incidence rate ratios of administrative sex (ref.: men) for deaths, stratified by quintiles of socio-economic position (SEP), for age groups 80+ (upper part), and 70-79 (lower part). Results for age groups 0-59 and 60-69 are not presented due to the absence of significant differences, attributed to a small number of events in these categories. Estimation was obtained using a Poisson regression model as no overdispersion was found in this triple interaction model.

The relationships between sex/gender, age, and SEP were investigated using a triple interaction term across these covariates. Findings revealed that, globally, for both age groups 70-79 and 80+, the sex/gender gap in mortality decreased with increasing SEP.

Supplementary Figure 9: Sensitivity analysis - incidence rate ratios of death notifications using the general population as denominator, employing negative binomial regression models with interaction terms

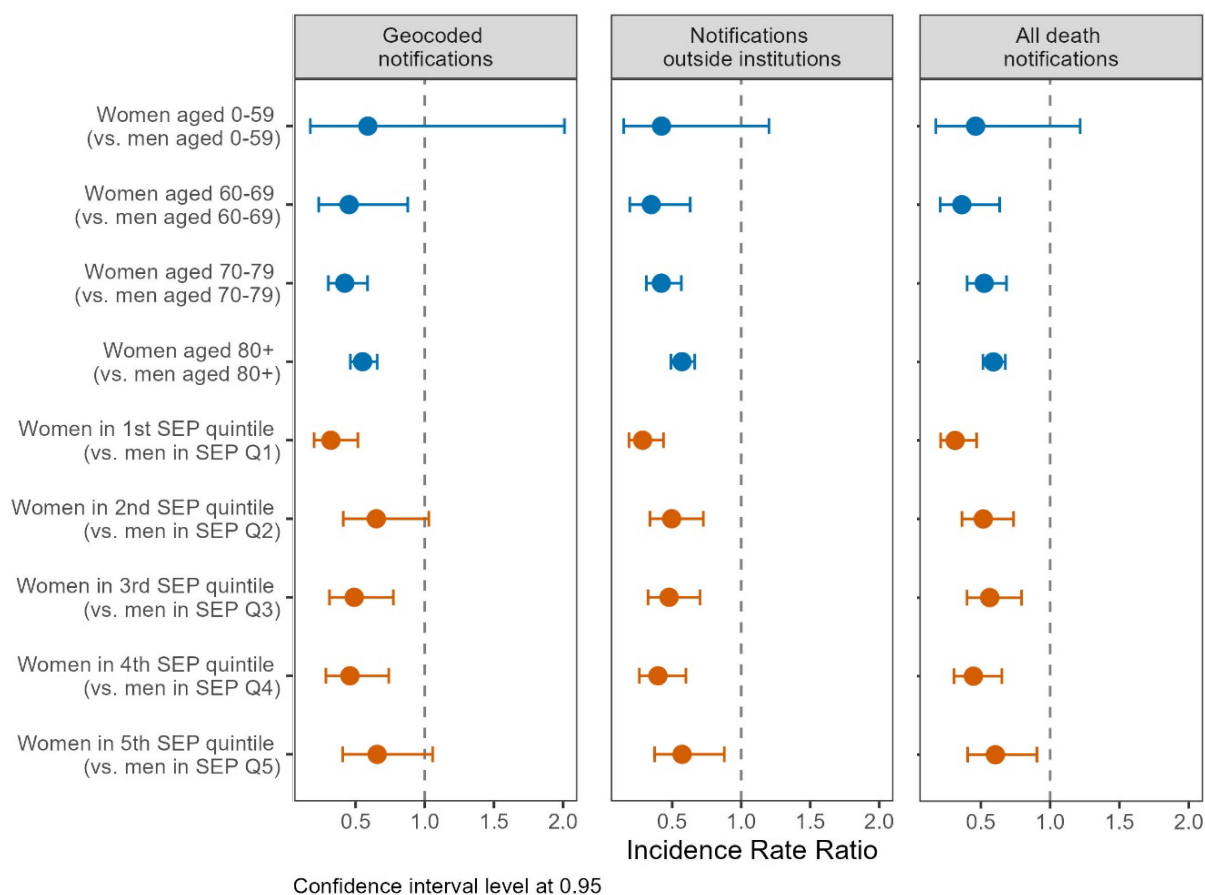


Figure S9 – Incidence rate ratios of administrative sex (ref.: men) for geocoded death notifications (column 1), notifications of death including institutional location (column 2) and all death notifications (column 3), stratified by age groups (upper part), and quintiles of socio-economic position (SEP, lower part), using general population as denominator, Canton of Vaud surveillance data 2020-2021, Switzerland

Notes: Geocoded death notifications (column 1), N=713 N; including institution-related addresses (column 2), N=968; including incomplete addresses without accurate retrievable SEP (column 3), =1175.

Coefficients of negative binomial models using interaction term between sex/gender and age and/or SEP quintiles reported in forest plots were estimated using the *contrast* function of the *emmeans*, for *estimated marginal means*, package in R (12). This function facilitates the pairwise comparisons by treating each category of age and SEP as the reference category, comparing women to men.

Reference

1. Office Fédéral de la Statistique. Registre fédéral des bâtiments et des logements (RegBL): Office Fédéral de la Statistique (2022) [Available from: <https://www.housing-stat.ch/fr/madd/public.html>].
2. Robinson D, Elias J. Package 'fuzzyjoin'. URL: <https://cran.r-project.org/web/packages/fuzzyjoin/fuzzyjoin.pdf>. (2018).
3. Pebesma EJ. Simple features for R: standardized support for spatial vector data. *R J.* (2018);10(1):439.
4. Population and Household Statistics Geodata (STATPOP GEOSTAT) (2020). Federal Office of Statistics; [Available from: <https://www.bfs.admin.ch/bfs/en/home/services/geostat/swiss-federal-statistics-geodata/population-buildings-dwellings-persons/population-housholds-from-2010.html>].
5. QGIS Development Team. QGIS Geographic Information System. Open Source Geospatial Foundation Project; (2022).
6. Statistique Vaud (2020). Population résidante permanente par âge exact_ sexe et origine_ Vaud_ 2017-2020 [Available from: <https://www.vd.ch/themes/etat-droit-finances/statistique/statistiques-par-domaine/01-population/etat-et-structure-de-la-population>].
7. Panczak R, Galobardes B, Voorpostel M, Spoerri A, Zwahlen M, Egger M. A Swiss neighbourhood index of socioeconomic position: development and association with mortality. *J Epidemiol Community Health.* (2012);66(12):1129-36.
8. Ihle A, Gabriel R, Oris M, Gouveia ÉR, Gouveia BR, Marques A, et al. Cognitive reserve mediates the relation between neighborhood socio-economic position and cognitive decline. *Dementia and Geriatric Cognitive Disorders Extra.* (2022);12(2):90-3.
9. Matthes KL, Zuberbuehler CA, Rohrmann S, Hartmann C, Siegrist M, Burnier M, et al. Selling, buying and eating—a synthesis study on dietary patterns across language regions in Switzerland. *British Food Journal.* (2021);124(5):1502-18.
10. Jeong A, Eze IC, Vienneau D, de Hoogh K, Keidel D, Rothe T, et al. Residential greenness-related DNA methylation changes. *Environ Int.* (2022);158:106945.
11. Panczak R, Berlin C, Voorpostel M, Zwahlen M, Egger M. The Swiss neighbourhood index of socioeconomic position: update and re-validation. *Swiss Medical Weekly.* (2023);153(1):40028.
12. Lenth RV. Emmeans: Estimated Marginal Means, aka Least-Squares Means (2023) [Available from: <https://CRAN.R-project.org/package=emmeans>].